# An Integral Projection Model for Giant Hogweed 

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I hereby declare that I am the sole author of this thesis. This is a true copy of the thesis, including any required final revisions, as accepted by my examiners.

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#### Abstract

In this thesis I investigate the impact that the choice of model may have on predictions of giant hogweed population dynamics. It has been shown in two case studies that population growth rates predicted by matrix population models may be biased if the number of plants sampled is low and plants are classified based on a continuous measure of their size [79]. These same studies have shown that integral projection models do not provide biased estimates of population growth rates for populations with few plants sampled. In chapter 2 I construct a density-independent integral projection model for giant hogweed population dynamics and I use it to verify that the population growth rates do not significantly differ from a previously published matrix model of giant hogweed population dynamics [39]. This research indicates that the conclusions made using the matrix population model are unaffected by the decision to discretize plant size.

A second major topic of research for giant hogweed populations is to predict the rate at which the species spreads. It has been shown that the rate of spread of a plant may be affected by the number of seeds produced by the population [68]. In chapter 3 I develop two density-dependent integral projection models for giant hogweed populations and compare the total seed production predicted by each model. In both models I allow for recruitment to be limited by competition among seedlings. However, in the second model I also allow the probability of flowering to depend on intraspecific density. I find that the model with density-dependent flowering has oscillatory seed production and that the model predicts significantly fewer seeds every other year compared to the model with density-independent flowering. I conclude that accounting for competition among adult plants may affect predictions of the rate of spread of giant hogweed.


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## Dedication

To my mom, Joan, and my dad, Peter

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## Chapter 1

## Literature Review

The spread of H. mantegazzianum is likely to continue through much of southern Canada over the next 25-100 yr with worsening ecological, economic, and health effects

Page et al., 2006 [69]

### 1.1 Background and motivation

Giant hogweed (Heracleum mantegazzianum Somm. \& Lev.) is a large, monocarpic perennial terrestrial plant native to the Caucasus mountain range [93]. The species is highly invasive, having successfully invaded 19 countries in Europe over the past 70 years, and is in the process of colonizing northeastern and northwestern North America [69]. As with many invasive, non-native plants, giant hogweed poses a serious threat to biodiversity in its invaded habitat [76]. Giant hogweed plants are able to suppress native plant species due to their large size and rapid sprouting in the spring [93] and due to their production of allelopathic chemicals that inhibit germination of competitors [42].

In addition to the impact on native flora, giant hogweed is a pest to humans. Giant hogweed plants produce phototoxic furanocoumarins as an insecticide, but these furanocoumarins will also cause third-degree burns or blindness in humans [93]. Finally, as an invasive riverside plant giant hogweed frequently blocks access to amenity areas [86]
and, during winter, plants die back leading to bare ground which makes the riverbank more susceptible to erosion [86]. For these reasons giant hogweed has been designated as a noxious weed in numerous jurisdictions, including the United States [1] and Ontario [2].

There has been more research on the European hogweed invasion since the invasion began earlier and is more advanced than the invasion in North America; however, giant hogweed invasion in North America remains a serious threat [69, 72, 90]. Although the species is unlikely to be eradicated in either Europe or North America, there is a silver lining in that much of what has been learned from the invasion of Europe may be applied to mitigate the effects of giant hogweed invasion in North America. In addition to in situ evaluations of giant hogweed control strategies, mathematical models of giant hogweed population dynamics have been constructed to better understand factors influencing giant hogweed invasion and control [38, 39, 60, 62, 63, 73, 72].

### 1.2 Matrix population models

### 1.2.1 Introduction to matrix models

For many plant species, including giant hogweed, the vital rates of individuals within a population differ with the age, stage, or size of the individual. To account for this heterogeneity, structured population models can be used which allow for the vital rates of individuals to be dependent on their state. Allowing vital rates to depend on state is particularly important for modelling invasive species since researchers can determine which individuals have the largest impact on population growth [18]. This knowledge may be used to develop effective management strategies.

By far the most common type of structured model used to describe plant population dynamics are matrix population models, with one recent study finding 355 out of 396 plant population modelling papers surveyed incorporating matrix models in some capacity [16]. Matrix models are discrete-time population models in which individuals are classified into a finite number of discrete classes [12]. The class of an individual is assumed to be representative of its state so that the vital rates of the individual are solely determined by the class it occupies.

To model the dynamics of a given population, matrix population models assume the number of individuals in each class in the next time-step will rely on how many individuals transition to that state. More specifically, matrix models assume the number of individuals in each class after one time-step may be written as a linear combination of the current
number of individuals in each class. The advantage of this approach is that simple matrix models are easy to construct, easy to parameterize, and methods to analyse matrix models are well developed.

For deterministic, density-independent matrix models of the form $n(t+1)=A n(t)$ the population vector, $n(t)$, will converge to a stable population distribution dependent on the demographic matrix, $A$ [12]. Once the population has reached its stable distribution, the population vector at the next time step will be a multiple of the current population vector. That is, $n(t+1)=\lambda n(t)$ for some value $\lambda$. The largest $\lambda$ that satisfies the equation is the dominant eigenvalue, which is the population growth rate in a matrix model. In addition, sensitivity and elasticity analysis may be performed to determine how small perturbations in the elements of $A$ affect population growth rate [12].

### 1.2.2 Previous matrix models for giant hogweed

Hüls and coauthors collected data from the Hesse region of Central Germany from 20022004 [38, 39, 71] while Pergl and coauthors collected data from the Slavkovský les region in the Western part of the Czech Republic from 2002-2005 [71, 72, 75]. Both sets of authors classified individuals into three stages of small, medium, or large vegetative plants along with a fourth stage for flowering plants. The classification made by Hüls and coauthors was dependent on plant height and the laminar width of the largest leaf while the classification made by Pergl and coauthors defined seedlings as small plants, juveniles as medium sized plants, and rosette plants as large plants. In both studies, the state of each plant was recorded every year which allowed the authors to determine the probability a plant in one class may transition to another state. Each research group then used the transition data to construct density-independent matrix models to predict the size and structure of the population in future years (Figure 1.1).

In addition to classification based on size, Hüls and coauthors classified each $2.5 \mathrm{~m}^{2}$ plot for which they collected data as either dense or open stands depending on the dominance of giant hogweed in the area [39]. The authors had originally hypothesized that open stands occur at the front of giant hogweed invasions and that these sparse stands are precursors to dense stands. However, the authors' analysis of their matrix model indicated that open stands do not have a higher intrinsic population growth rate than dense stands. They concluded that dense stands are saturated with large individuals which suppress germination while open stands are likely limited by irregular disturbances such as by mowing or grazing. Furthermore, the authors found that population growth rate in dense stands is more sensitive to survival of existing plants while population growth rate is more sensitive to the growth of individuals in open stands.


Figure 1.1: Life cycle diagram of the matrix models developed by Hüls and coauthors [38] and by Pergl and coauthors [73]. Nodes pictured represent (S) small vegetative plants, (M) medium vegetative plants, (L) large vegetative plants, and (R) reproductive plants. Dashed lines indicate reproduction. Arrows marked with an asterisk $\left(^{*}\right)$ indicate transitions that were included in the model developed by Hüls and coauthors but not the model developed by Pergl and coauthors. Retrogression is possible due to grazing or as a response to stress [85].

For 2002-2003, Hüls and coauthors found that populations were in decline [39]. The authors attributed this decline to the unusually hot and dry conditions of the summer of 2003. These conditions led to low productivity which resulted in increased mortality. They found that plants responded to this increased mortality by increasing reproductive output to take advantage of the decrease in ground cover. Pergl and coauthors [71] compared their results to those of Hüls and coauthors and found that their surveyed populations are decreasing in the Czech Republic, complementing the conclusion by Hüls and coauthors that populations are decreasing in Germany. However, the German and Czech study periods overlapped and thus the decrease in populations in the Czech Republic may also be due to the extreme temperature during the study period.

### 1.3 Integral projection models

### 1.3.1 Why an integral projection model?

A major drawback of matrix projection models, such as those developed by Hüls and coauthors [38, 39, 71] and Pergl and coauthors [71, 72, 73], is that they require modellers to classify plants into discrete stages in their life-cycle. In a plant such as giant hogweed, the divisions between classes are often based on continuous variables. The resulting classes are constructed arbitrarily with little regard for the underlying biology of the plant. For example, in the matrix model by Hüls and coauthors stage 1 plants were defined as having laminar width between $3.5-20 \mathrm{~cm}$ and height less than 50 cm while stage 2 plants have laminar width between $20-50 \mathrm{~cm}$ and height greater than 50 cm but not reaching the canopy. The matrix models cannot account for variation in vital rates for plants within each class. As a result, the choice of boundaries between classes may have an impact on model predictions [79]. Additionally, the number of classes has been shown to bias predicted population growth rate, and population growth rate may be underestimated when there are not very many individuals [79].

There are several continuous size individual-based models (IBMs) for giant hogweed [96, 63, 60, 61, 62]. The hypotheses in this thesis may be answered using IBMs similar to previously published IBMs; however, they do not offer any advantages over integral projection models. Furthermore, the hypotheses in this thesis relate to previously published matrix models. I opt to use integral projection models throughout this thesis due to their structural similarity to matrix models.

### 1.3.2 Overview of integral projection models

In his 1998 dissertation Easterling introduced integral projection models (IPMs) [25]. These models are discrete-time, size-structured population models that allow for individuals within a population to be described by a continuous trait using an integral operator [25] (as cited in [24]). This type of model allows researchers to avoid arbitrary classifications of individuals based on size by allowing vital rates in the population to depend on a continuous measure of size. For example, we may use the height of a plant as the relevant trait, in which case the population may be represented by a size distribution $n(x, t)$ where $x$ is the height of an individual, $t$ is the time, and $n(\cdot, t)$ is the distribution of individuals at time $t$. Projection from one time-step to the next relies on a kernel, $K(y, x)$, that yields the probability that a plant will transition from state $x$ to state $y$ over a timestep. This kernel is comprised of various size-dependent vital rates functions that may be determined through regression. The expected population at a given size in the next time step is dependent on the contribution of every plant that may transition to that size. Thus we integrate the kernel multiplied by the size distribution at time $t$ over the size $x$ to get the projected size distribution at the next time-step. Mathematically, this is written as:

$$
\begin{equation*}
n(y, t+1)=\int_{\omega} K(y, x) n(x, t) d x \tag{1.1}
\end{equation*}
$$

The form of the kernel will be determined by the life cycle of the species and parametrization is then an exercise in statistical regression to fit survival, growth, and reproduction components of the kernel. The survival and reproduction functions are typically found using logistic regression while the growth function may be found using linear or non-linear regression. Since being introduced by Easterling, IPMs have been developed to accommodate discrete states, time lags, spatial structure, environmental stochasticity, and demographic stochasticity [26]. Additionally, in making the adjustment from matrix models to IPMs, the potential for analysis is not greatly diminished. The theory used to analyze IPMs is slightly less developed; however, most of the useful properties of matrix models such as the population growth rate, stable size distribution, and net reproductive ratio may still be computed for IPMs [24, 26].

### 1.3.3 Density-dependent integral projection models

The probability of survival, probability of flowering, expected increase in plant size, and expected number of recruits are all functions that may depend on the presence of other plants. Accounting for the competition among giant hogweed plants will therefore be necessary when modelling the course of giant hogweed invasion since plant density will change over time. However, there are limitations to the construction of density-dependent IPMs.

As is necessary for density-dependent matrix models, the forms of the vital rate functions must be restricted in order to parameterize the model. However, density-dependent IPMs have an advantage over matrix models in that they may be parameterized in cases where matrix models cannot. In particular, in matrix models each element may depend on giant hogweed density, but not every transition will have enough observations for such a function to be parameterized. This limitation is overcome in integral projection models by assuming that population dynamics are describes using a continuous kernel and assuming that the functions describing the kernel may be fit using regression.

### 1.4 Outline of thesis

This thesis is divided into two main chapters. In the first chapter I develop a densityindependent integral projection model analogous to the matrix model constructed by Hüls and coauthors [39]. I compare the predicted population growth rates for each model to determine if the conclusions originally reached by Hüls and coauthors are affected by their use of a matrix model, as implied by the results of Ramula and coauthors [79]. In the second chapter I develop a density-dependent integral projection model in order to examine the transient and asymptotic behaviour of giant hogweed populations. I use this model to determine the impact that competition among adult plants will have on seed production in giant hogweed stands. The two chapters are united in that they investigate how increasing model complexity affect our understanding of giant hogweed spread.

## Chapter 2

# An Integral Projection Model for Giant Hogweed Management Decisions 

Les fleurs sont si contradictoires!<br>Antoine de Saint-Exupéry,<br>Le Petit Prince

### 2.1 Introduction

A fundamental concern in the management of invasive species is to assess the rate at which invasive population size changes. For many populations, researchers construct matrix models in which individuals are classified into discrete stages based on the individual's state and use these models to determine long-term population growth rates [12]. However, the population growth rate predicted by a matrix model may be biased if this classification is based on a continuous measure of size and the sample size is small [20, 79]. This potential for bias in estimated population growth rate may affect management priorities or may lead officials to underestimate the threat of an invasive species.

Giant hogweed (Heracleum mantegazzianum Somm. \& Lev.) is an invasive species that has become naturalized in parts of Europe and North America [69]. Giant hogweed plants pose a threat to biodiversity in its invaded ranges [76], can negatively affect human
health [93], and facilitate riverside erosion [86]. These potential impacts have prompted researchers to evaluate control strategies for the species. For example, Hüls and coauthors [39] investigated the relation between populations in dense, monospecific stands near the centre of infestations and populations in open stands near the margins that were intermixed with surrounding vegetation. The authors collected demographic data from the two types of giant hogweed stands and used the data to parameterize matrix models for the populations [39]. The authors noted that if the population growth rate in open stands was higher than the population growth rate in dense stands then they may represent the front of an invasion and should be prioritized for control. Using their matrix model, Hüls and coauthors found no significant difference in population growth rates between the stand types. Since open stands do not represent the front of the invasion, the authors concluded that they should not be prioritized for control. However, the use of a stage-structured model for giant hogweed populations may be unsuitable to address the authors' hypothesis. In particular, the low sample size of the open stand data set may produce a biased estimate of the population growth rate [20, 79], which in turn may affect the conclusions made by Hüls and coauthors.

Integral projection models (IPMs) are an alternative to matrix models that avoid discretizing populations that have continuous traits such as size [24]. Previous authors have speculated that avoiding this discretization would prevent bias in the predicted population growth rate since plants vital rate functions would be more precise [27]. Ramula and coauthors tested the impact that sample size has on matrix models and IPMs and found that for large demographic data sets IPMs yield the same population growth rate as matrix models [79]. As sample size decreased, the authors found that the mean estimate of population growth rate predicted by the IPMs remained the same; however, the mean estimate of population growth rate predicted by the matrix models changed. Ramula and coauthors concluded that matrix models may yield biased predictions of population growth rate when the data set used to parameterize the model is small. For both of the species studied by Ramula and coauthors, the population growth rate predicted by the matrix model decreased once the sampled population dropped below 300 plants.

The results of Hüls and coauthors may be affected since the data sets used to parameterize their matrix model vary in size [39]. The open stand data set is small enough to produce biased estimates of population growth rate. In contrast, the dense stand data set should yield an unbiased estimate of population growth rate. The conclusions of Hüls and coauthors that depend on these predicted population growth rates could therefore change with the use of an integral projection model. This possibility that the selection of the modelling approach may affect management decisions is my primary motivation for revisiting the data set collected by Hüls and coauthors.

In this chapter I investigate whether the decision to use a matrix model for giant hogweed population dynamics affected the conclusions made by Hüls and coauthors [39]. The results of Ramula and coauthors imply that the demographic data set used to parameterize the matrix model for dense stands is large enough to produce an unbiased estimate of population growth rate [79]. However, the data set of plants in open stands is too small to guarantee that the estimate of population growth rate is unbiased. To test this hypothesis I constructed an integral projection model and parameterized it for both stand types using data collected by Hüls and coauthors during the 2002-2003 growing season. I then used the population growth rates predicted by the integral projection model to revisit the conclusion made by Hüls and coauthors that open stands are not precursors to dense stands. My analysis will demonstrate whether the choice between using a matrix model or integral projection model for populations with a low number of records may affect management recommendations.

### 2.2 Methods

### 2.2.1 Study species

Heracleum mantegazzianum is a monocarpic perennial species in the family Apiaceae [69] that propagates exclusively by seed [78]. Seeds undergo dormancy breaking via wet and cold stratification [58] and typically germinate in spring the year after they set, but under suitable conditions may germinate the same year in autumn [93]. In the year a plant germinates it will direct much of its resources towards its taproot [69], causing the accumulation of aboveground biomass to be slow the first year [93]. Growth is much more vigorous in subsequent years once the root system is developed. After 3-5 years of growth plants will have accumulated enough resources to begin flowering [69].

If root reserves are sufficient at the beginning of the growing season a plant will initiate flowering that year [69]. Plants that will flower begin the year with vigorous vegetative growth, followed by stem elongation and umbel formation sometime between late April and early June [10, 93]. This vegetative growth and stem elongation causes flowering plants to grow extremely large, with most flowering plants between 3 m and 4 m tall and some recorded up to 5.5 m in height [78]. Flowering typically occurs over 5 or 6 weeks in early-mid summer while seeds ripen later in the summer in August and September [69].

Occurance of $H$. mantegazzianum in its invaded range may be limited to sparse stands of only a few individuals, linear stands along dispersal corridors, or dominant stands in
which the species attains nearly $100 \%$ ground cover [39]. Plants have an affinity for sites that have been disturbed by human activity or flooding. Once established, plants suppress competing species and promote the growth of conspecifics [39].

### 2.2.2 Experimental data collection

I used the data set collected in Germany by Hüls and coauthors to parameterize my model [39]. I opted to use these data since the difference in the size of the open stand data set and the size of the dense stand data set may be large enough to influence the conclusions of the authors' original paper [20, 79]. Annual transition data for vegetative plants were collected from seventy-six plots located in five sites in Germany between 2002 and 2004 by Dr. Jörg Hüls and coauthors [38, 39]. Sites were located in Allendorf, Druseltal, Frankenberg (Burgwald), Kassel (Dönche), and Viermünden (Table 2.1). I ultimately did not use data collected from Allendorf since the site was disturbed by cattle. Measurements of height, the petiole diameter for the largest leaf, and laminar width of the largest leaf were recorded in 2002 and 2003 for all plants that were 10 cm in height or taller. Data were also collected on individuals in 2004; however, since the data set was originally intended to parameterize a matrix model, plants were classified on site without recording morphological data during data collection in 2004. As a result, I used only the 2002-2003 data sets collected by Hüls and coauthors to parameterize the model.

Table 2.1: Site information for data collected by Hüls and coauthors [39]. The maximum distance between sites is 56 km .

| Site | Position |  | Elevation | Area | Number of <br> Plots | Number of <br> Individuals |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Lat $(N)$ | Lon $(E)$ |  |  | 16 | 216 |
| Burgwald | $51^{\circ} 01^{\prime} 37^{\prime \prime}$ | $8^{\circ} 45^{\prime} 03^{\prime \prime}$ | 330 m | 600 | 16 | 126 |
| Dönche | $51^{\circ} 17^{\prime} 50^{\prime \prime}$ | $9^{\circ} 25^{\prime} 22^{\prime \prime}$ | 290 m | 4500 | 18 | 27 |
| Druseltal | $51^{\circ} 18^{\prime} 15^{\prime \prime}$ | $9^{\circ} 24^{\prime} 49^{\prime \prime}$ | 335 m | 30 | 10 | 216 |
| Viermünden | $51^{\circ} 05^{\prime} 28^{\prime \prime}$ | $8^{\circ} 49^{\prime} 41^{\prime \prime}$ | 315 m | 1400 | 16 | 216 |

### 2.2.3 Model formulation

The matrix model of Hüls and coauthors consisted of four classes in total: three for vegetative plants classified by size and one for reproductive plants [39]. In order to be able to
compare my model predictions to the prediction made by Hüls and coauthors, I first constructed an integral projection model as similar as possible to the original matrix model. Instead of discretizing vegetative plants based on size I used plant height ( cm ) as a continuous measure of plant size. Since the species is monocarpic I did not include a separate class for reproductive plants.

If we let $n(x, t)$ be the size distribution of plants at time $t$ with lower and upper bounds $L$ and $U$, respectively, then the population distribution at the next time step will be governed by the integral equation [24]:

$$
\begin{equation*}
n(y, t+1)=\int_{L}^{U} K(y, x) n(x, t) \mathrm{d} x \tag{2.1}
\end{equation*}
$$

for an appropriate kernel $K(y, x)$. The kernel may be decomposed into the sum of survivalgrowth $(P(y, x))$ and fecundity $(F(y, x))$ parts so that [24]:

$$
\begin{equation*}
K(y, x)=P(y, x)+F(y, x) . \tag{2.2}
\end{equation*}
$$

The survival-growth kernel is the product of the probability a plant of size $x$ is still a vegetative plant the next year multiplied by the probability the plant will grow to size $y$. The probability a plant remains vegetative may itself be decomposed into the product of surviving to the next year, $p_{s}(x)$, and the probability the plant does not flower, $1-p_{r}(x)$, where $p_{r}(x)$ is the probability a plant of size $x$ flowers [83]. I defined the growth kernel, $G(y, x)$, to be the conditional probability that a plant of height $x$ that remains vegetative will have height $y$ in the next year. Thus the survival-growth kernel may be written as [83]:

$$
\begin{equation*}
P(y, x)=p_{s}(x)\left[1-p_{r}(x)\right] G(y, x) . \tag{2.3}
\end{equation*}
$$

The fecundity kernel is the product of the expected number of new plants that survive until the annual census multiplied by their expected size distribution [52]. I defined recruits to be plants that germinated and established earlier in the year of the census that have at least one leaf with a laminar width of 3.5 cm or greater. First-year plants that did not have a leaf with a laminar width of at least 3.5 cm were not included in the census since their survival to the following year was negligible [39]. The expected number of recruits is the
product of the probability a plant of size $x$ survives one more year and the probability that the plant reproduces in that year, $p_{s}(x) p_{r}(x)$, multiplied by the expected number of recruits produced by one flowering plant, $f_{e}$. I assumed mixing-at-birth for recruit size distribution [13], meaning that the size of the parent plant does not affect recruit size. Under this assumption I denoted recruit size distribution by $c_{0}(y)$. Altogether, the fecundity kernel may be written as [52]:

$$
\begin{equation*}
F(y, x)=p_{s}(x) p_{r}(x) f_{e} c_{0}(y) \tag{2.4}
\end{equation*}
$$

Thus population dynamics will be described entirely by the integral equation:

$$
\begin{equation*}
n(y, t+1)=\int_{L}^{U} p_{s}(x)\left[\left(1-p_{r}(x)\right) G(y, x)+p_{r}(x) f_{e} c_{0}(y)\right] n(x, t) \mathrm{d} x \tag{2.5}
\end{equation*}
$$

### 2.2.4 Parameterization

## Bounds on plant size

The lower and upper bounds of plant size, $L$ and $U$, must be chosen carefully in integral projection models in order to minimize the number of plants that are needlessly excluded by the model. This phenomenon, known as unintentional eviction, occurs when plants are projected to have a size smaller than the lower bound or larger than the upper bound [27, 101]. Since only plants within the bounds are included in the population vector, plants that are projected to leave the bounds are "evicted" from the population and no longer have an impact on population dynamics.

The lower and upper bounds were initially selected to be the heights of the shortest $(10 \mathrm{~cm})$ and tallest ( 228 cm ) plants, respectively. However, plants may be slightly shorter or taller than these bounds and so I extend the bounds to $L=5 \mathrm{~cm}$ and $U=250 \mathrm{~cm}$. I used the process outlined by Williams and coauthors to determine if unintentional eviction affects my results [101].

In my model eviction may occur either in the fecundity kernel or the survival-growth kernel. Eviction through the fecundity kernel was prevented numerically by truncating and normalizing the recruit size distribution, $c_{0}(y)$. Unfortunately, eviction through the survival-growth kernel is more difficult to mitigate.

The unconditional probability of eviction through the survival-growth kernel is a function defined to be $\rho(x)=p_{s}(x)\left[1-\int_{L}^{U} G(y, x) \mathrm{d} y\right]$. A maximum of $0.016 \%$ of plants may be evicted in open stands while a maximum of $10.3 \%$ of plants may be evicted in dense stands. However, the maximum percent of evicted plants is not useful without accounting for the size distribution of plants.

A more meaningful measure of eviction is to take the inner product between the unconditional probability of eviction and the stable size distribution. I first break down the unconditional probability of eviction into the probability of eviction through the lower bound, $\rho_{L}(x)=p_{s}(x) \int_{-\infty}^{L} G(y, x) \mathrm{d} y$, and the probability of eviction through the upper bound, $\rho_{U}(x)=p_{s}(x) \int_{U}^{\infty} G(y, x) \mathrm{d} y$. I found the inner product between the probabilities of eviction with the stable size distribution to determine the proportion of plants that are evicted through each bound at the stable size distribution. For open stands, I found that $0.00022 \%$ of plants are evicted through the lower bound while $0.00000051 \%$ of plants are evicted through the upper bound. For dense stands, I found that $2.13 \%$ of plants are evicted through the lower bound and $0.041 \%$ of plants are evicted through the upper bound.

I then determined the rate at which eviction will decrease as the bounds change. For open stands, the rate at which eviction decreases as the lower bound decreases is $\mathrm{d} \lambda_{L}=$ $9.3 \times 10^{-7}$ and the rate at which eviction decreases as the upper bound increases is $\mathrm{d} \lambda_{U}=$ $4.6 \times 10^{-9}$. For dense stands, the corresponding rates of change are $\mathrm{d} \lambda_{L}=0.00845$ and $\mathrm{d} \lambda_{U}=0.000441$. None of these values are particularly concerning except for $\mathrm{d} \lambda_{L}$ for open stands.

In order to determine if eviction through the lower bound in dense stands is a problem, I lowered $L$ to the minimum physically possible height of 0 cm and recomputed the population growth rate. Since the population growth rates were identical up to two significant figures, I concluded that a minimum bound of $L=5 \mathrm{~cm}$ is reasonable to use in the model.

## Probability of survival and reproduction

Both the probability of survival and the probability of flowering for the data collected by Hüls and coauthors [39] were fit using logistic regression with height as the explanatory variable (Figure 2.1). For direct comparison to the predictions of the matrix model, the probability of survival and probability of flowering were fit once using only plants from open stands then again using only plants from dense stands.


Figure 2.1: Data and best-fit functions for relationships between plant height and vital rates for open and dense stands. Survival probability fit using logistic regression (a, b), probability of reproduction fit using logistic regression ( $\mathrm{c}, \mathrm{d}$ ), and expected height of surviving plants fit using linear regression (e, f) in open (a, c, e), and dense (b, d, f) stands of giant hogweed in the Hesse region of Germany. Ticks on the figures fit using logistic regression represent data points with plant height indicated. Plants that survive ( $\mathrm{a}, \mathrm{b}$ ) and plants that flower (c, d) have ticks on the top while plants that do not survive (a, b) or do not flower (c, d) have ticks on the bottom. All data collected by Hüls and coauthors [39].

## Inter-annual growth

Inter-annual growth was fit for both open and dense stands using linear regression (Figure 2.1). I also fit inter-annual growth using several nonlinear parametric functions and compared these fits to the linear model using the Akaike information criterion (AIC; Table 2.3) [9]. Since weak nonlinearity has been previously shown to significantly affect predicted population growth rate [19], I tested to see if any of the nonlinear models with a lower AIC value yielded different values for predicted population growth rate. In addition to the parametric nonlinear models, I fit a generalized additive model using the gam function from the mgcv package in R [102]. The generalized additive model was fit using default parameters and the results of the IPM that used the GAM growth function was compared to the results of the IPM that assumed the growth function was linear [82].

The suitability of the linear regression was evaluated by testing if the data is homoskedastic and whether residuals are normally distributed. I used the Breusch-Pagan test to test for heteroskedasticity and the Shapiro-Wilk test to check if residuals are normally distributed. Furthermore, in order to determine the impact that nonnormal variance in residuals may have on estimated population growth rate I fit both a linear model and a generalized additive model that each had their error distribution determined using kernel density estimation [82].

## Recruitment

The expected number of recruits, $f_{e}$, was determined using the same method as Hüls and coauthors [39]. In order to find the expected number of recruits I determined the total number of new plants in each stand type and divided by the total number of flowering plants in the respective stand type. Instead of categorizing these plants by stage as was done in the matrix model, recruit size was assumed to follow a log-normal distribution; however, in order to prevent unintentional eviction I used a truncated log-normal distribution in the numerical implementation. The validity of assuming a log-normal distribution was tested by log-transforming the height of each recruit and using the Shapiro-Wilk test for normality.

## Outliers

One plant was removed from the analysis of open stands since its height in 2003 was clearly recorded incorrectly. The plant reportedly decreased in size from 85 cm to 7 cm , yet the

Table 2.2: Vital rate function forms and parameter estimates for the integral projection model. Standard errors for parameter estimates are in parentheses. All data collected by Hüls and coauthors [39].

| Function | Stand type | Fitted function |
| :--- | :--- | ---: |
| Survival $\left(p_{s}\right)$ | Open | $\operatorname{Logit}\left(p_{s}(z)\right)=-0.845_{(0.873)}+0.041_{(0.014)} z$ |
|  | Dense | $\operatorname{Logit}\left(p_{s}(z)\right)=-0.251_{(0.241)}+0.021_{(0.004)} z$ |
| Reproduction $\left(p_{r}\right)$ | Open | $\operatorname{Logit}\left(p_{r}(z)\right)=-2.645_{(0.973)}+0.036_{(0.012)} z$ |
|  | Dense | $\operatorname{Logit}\left(p_{r}(z)\right)=-4.156_{(0.451)}+0.032_{(0.004)} z$ |
| Growth $(G)$ | Open | $G\left(z_{1}, z\right)=75.96_{(12.64)}+0.38_{(0.18)} z+\mathcal{N}\left(0,22.15^{2}\right)$ |
|  | Dense | $G\left(z_{1}, z\right)=32.70_{(5.25)}+0.68_{(0.07)} z+\mathcal{N}\left(0,38.40^{2}\right)$ |
| Mean recruits $\left(f_{e}\right)$ | Open | $f_{e}=0.62$ |
|  | Dense | $f_{e}=0.92$ |
| Recruit size $\left(c_{0}\right)$ | Open | Dense |

plant remained a stage 2 individual ( $>50 \mathrm{~cm}$ ) with a laminar width of 61 cm for its largest leaf. In addition to this record, one outlier was removed from the analysis of dense stands since it had an unusually large height increase from 32 cm to 228 cm in one year. When growth was fit the standard deviation of the residuals was 37 cm while the residual for this point was 155 cm . Since its growth was more than four standard deviations larger than the mean increase I concluded height was recorded incorrectly for this plant as well and removed it from the data set for dense stands.

### 2.2.5 Life cycle analysis

The population growth rate for each model was determined by finding the dominant eigenvalue of the demographic kernel, defined to be the largest value of $\lambda$ that satisfies [24]:

$$
\begin{equation*}
\lambda w(y)=\int_{L}^{U} K(y, x) w(x) \mathrm{d} x \tag{2.6}
\end{equation*}
$$

for some function $w(x)$ called the right eigenvector. The eigenvalue and right eigenvector were computed numerically by first using the midpoint rule to produce a matrix equation,
$n(t+1)=A n(t)$, to approximate the IPM. The code for this approximation was extensively modified from code provided by Rees and coauthors [82] and may be found on UWSpace. I then found the largest $\lambda$ and its corresponding right eigenvector $w$ that satisfy the equation $\lambda w=A w$.

A unique population growth rate and corresponding right eigenvector are guaranteed since the kernel is power positive [27]. The dominant right eigenvector of the IPM represents the stable size distribution of the population which is the size distribution the population will converge to over time. In addition to the right eigenvector, there is a left eigenvector corresponding to the dominant eigenvalue that satisfies:

$$
\begin{equation*}
\lambda v(x)=\int_{L}^{U} K(y, x) v(y) \mathrm{d} y \tag{2.7}
\end{equation*}
$$

The dominant left eigenvector is also guaranteed to exist since the kernel is power positive and the dominant left eigenvector provides the relative reproductive values for the population [26]. The relative reproductive values are a measure of how a plant of a given size will impact future population size.

### 2.2.6 Bootstrap confidence intervals

I used bootstrapping to compare between different values for population growth rate and to compare mean observed plant height to the mean height predicted by the IPM. For population growth rate and mean predicted height, I computed 5,000 bootstrap samples for both the open stand data set and the dense stand data set. Each bootstrap sample was used to reparameterize the integral projection model, which was then used to determine the bootstrap statistics, $\bar{h}_{b}$ and $\lambda_{b}$. The bootstrap $95 \%$ confidence intervals were given by the $2.5 \%$ and $97.5 \%$ quantiles of the bootstrap distribution. It should be noted that any bootstrap sample that had two or fewer recruits were discarded and another sample drawn in its place. This requirement was necessary to ensure the recruit size distribution could be determined for the IPM. Bootstrapping for mean observed height was performed at the same time as population growth rate.

### 2.3 Results

### 2.3.1 Results of statistical analysis

## Inter-annual growth

I found that the linear model yielded the second lowest AIC value for the dense stand data set and the fifth lowest AIC value for the open stand data set. However, the differences in AIC values between the linear model and nonlinear models were small enough to not be of much concern $(\Delta \mathrm{AIC}<4$; Table 2.3). Furthermore, I found that the estimates of population growth rate provided by the nonlinear models were identical to the estimates of population growth rate provided by the linear model up to two significant figures. Since the potential presence of nonlinearity did not affect population growth rate and the quality of the fit (as measured by AIC) was only slightly improved by including nonlinearity I decided to use the linear model for growth.

Table 2.3: Evaluation of the linear and nonlinear functions used to model inter-annual growth of giant hogweed in Germany [39]. The best performing fit for each data set is in bold.

| Form | df | AIC |  |
| :--- | :---: | :---: | :---: |
|  |  | Open stands | Dense stands |
| $y=a x+b$ | 2 | 302 | 2416 |
| $y=a x^{2}+b x$ | 2 | 303 | 2425 |
| $y=a x^{2}+b x+c$ | 3 | 302 | 2417 |
| $y=a x e^{-b x}$ | 2 | 299 | 2421 |
| $y=a\left(1-e^{-b x}\right)$ | 2 | $\mathbf{2 9 8}$ | 2420 |
| $y=a x /(b+x)$ | 2 | 299 | 2418 |
| $y=a x^{b}$ | 2 | 300 | $\mathbf{2 4 1 5}$ |

The Breusch-Pagan test failed to reject the hypothesis that the data is homoskedastic in both open stands ( $p=0.23$ ) and dense stands $(p=0.074)$. The Shapiro-Wilk test failed to reject the hypothesis that residuals are normally distributed for open stands ( $p=0.84$ ); however, the test does reject the hypothesis that residuals are normally distributed for dense stands $(p=0.016)$. The predicted population growth rate for populations in dense stands was determined in three ways: using linear regression for growth with normally distributed errors, using linear regression for growth with the error distribution found using kernel
density estimation, and using a generalized additive model with the error distribution found using kernel density estimation. The predicted population growth rates found using each of the three methods were identical up to two significant figures. Therefore, I concluded that the growth kernel may be represented by a Gaussian conditional probability distribution with mean scaling linearly with height and constant variance.

## Recruitment

I log-transformed the heights of all seedlings and tested the normality of the distribution using the Shapiro-Wilk test. The Shapiro-Wilk test yielded $p=0.68$ for recruits in open stands, $p=0.77$ for recruits in dense stands, and $p=0.58$ for recruits in the combined data set. These results indicate that I cannot reject the hypothesis that recruit size follows a log-normal distribution. However, I note that the high p-value for open stands may be due to the low sample size which makes the null hypothesis unlikely to be rejected. A histogram of log-transformed recruit height in open stands indicates the distribution may be uniform (Figure 2.2); however, since the sample size is low and the recruit size distribution in dense stands appears to be log-normal (Figure 2.2), I assumed that the distribution in open stands is log-normal as well.



Figure 2.2: Histogram of the observed distribution of recruit size and fitted log-normal distribution for open $(\mathrm{n}=8, \mathrm{a})$ and dense $(\mathrm{n}=67, \mathrm{~b})$ stands of giant hogweed in the Hesse region of Germany. All data collected by Hüls and coauthors [39].

### 2.3.2 Results of demographic analysis

The population growth rates I calculated from the integral projection model are much less than one, like those found by Hüls and coauthors using a matrix model [39], and indicate that populations in both open and dense stands are in decline (Table 2.4). Similar to the results of the matrix model, the bootstrap confidence intervals for population growth rate in open and dense stands overlap. This overlap indicates that the difference in population growth rates between the two stand types is not statistically significant.

Table 2.4: Population growth rate $(\lambda)$, bootstrap estimate of population growth rate $\left(\lambda_{b}\right)$, and bootstrap $95 \%$ confidence interval (CI-, CI+). Bootstrap values were computed using 5,000 bootstrap samples. All data collected by Hüls and coauthors [39].

| Stand type | Model | N | $\lambda$ | $\lambda_{b}$ | $($ CI-, CI +$)$ |
| :--- | :--- | :---: | :---: | :---: | :---: |
| Open | Matrix | 103 | 0.76 | 0.77 | $(0.55,1.05)$ |
|  | IPM | 103 | 0.69 | 0.72 | $(0.45,1.13)$ |
| Dense | Matrix | 554 | 0.75 | 0.75 | $(0.70,0.81)$ |
|  | IPM | 554 | 0.73 | 0.73 | $(0.67,0.79)$ |

The large confidence interval for open stands in consistent with the results of Ramula and coauthors [79]. To determine if sample size affects my results, I performed bootstrapping with larger sample sizes for each replicate (Table 2.5). As sample size increased, the size of the confidence intervals decreased. However, even with the increased sample size I conclude that there is no significant difference in population growth rate between open and dense stands.

The stable size distributions for open and dense stands were computed and compared to the observed size distributions (Figure 2.3). The mean values for plant height from the stable size distributions are 86 cm for open stands and 77 cm for dense stands. In contrast, the mean heights determined from data recorded in 2003 are 95 cm for open stands and 73 cm for dense stands. The bootstrap confidence intervals for the predicted and observed values of height overlap in both open and dense stands (Table 2.6). Therefore, the mean heights at stable size distribution are not significantly different than their corresponding observed heights. These results are in contrast to the matrix model which determined that the mean predicted height for open stands is significantly higher than the mean observed height [39].

Table 2.5: Population growth rate $(\lambda)$, bootstrap estimate of population growth rate $\left(\lambda_{b}\right)$, and bootstrap $95 \%$ confidence interval (CI-, CI+) as the bootstrap replicate sample size $(N)$ increases. Bootstrap values were computed using 5,000 bootstrap samples. All data collected by Hüls and coauthors [39].

| Stand type | N | $\lambda$ | $\lambda_{b}$ | $($ CI-, CI +$)$ |
| :--- | :---: | :---: | :---: | :---: |
| Open | 100 | 0.69 | 0.72 | $(0.44,1.15)$ |
|  | 200 | 0.69 | 0.70 | $(0.50,0.96)$ |
|  | 300 | 0.69 | 0.70 | $(0.53,0.91)$ |
|  | 400 | 0.69 | 0.70 | $(0.55,0.88)$ |
|  | 500 | 0.69 | 0.69 | $(0.56,0.85)$ |
|  | 600 | 0.69 | 0.69 | $(0.57,0.84)$ |
|  | 100 | 0.73 | 0.74 | $(0.59,0.92)$ |
|  | 200 | 0.73 | 0.73 | $(0.63,0.86)$ |
|  | 300 | 0.73 | 0.73 | $(0.65,0.83)$ |
|  | 400 | 0.73 | 0.73 | $(0.66,0.81)$ |
|  | 500 | 0.73 | 0.73 | $(0.67,0.81)$ |
|  | 600 | 0.73 | 0.73 | $(0.67,0.79)$ |




Figure 2.3: Histogram of observed size distribution in 2003 vs predicted stable size distribution for open (a) and dense (b) stands. All data collected by Hüls and coauthors [39].

Table 2.6: Mean plant height $(\bar{h})$, bootstrap mean plant height $\left(\bar{h}_{b}\right)$, and bootstrap $95 \%$ confidence intervals (CI-, CI+). Bootstrap values were computed using 5,000 bootstrap samples. All data collected by Hüls and coauthors [39].

| Stand type | Source | $\bar{h}$ | $\bar{h}_{b}$ | $($ CI-, CI +$)$ |
| :--- | :--- | :---: | :---: | :---: |
| Open | Observed | 95 | 95 | $(87,103)$ |
|  | Predicted | 86 | 87 | $(76,98)$ |
| Dense | Observed | 73 | 73 | $(68,78)$ |
|  | Predicted | 77 | 78 | $(73,84)$ |

### 2.3.3 Impact of outliers

I re-parameterized the integral projection model using both the open and dense stand data sets without outliers to see how the exclusion of these plants affect model predictions. In both cases the population growth rate was unaffected by the removal of outliers up to two significant figures.

### 2.4 Discussion

In this chapter I developed a density-independent integral projection model for giant hogweed and parameterized it using two different data sets collected by Hüls and coauthors [39]. I found that the decision to categorize plants into discrete stages based on size had no significant impact on the conclusions of Hüls and coauthors and therefore did not affect the authors' suggested management priorities. This result does not support the hypothesis that model discretization significantly affects predicted population growth rate [20, 79].

Several authors have concluded that matrix models will not be biased for large demographic data sets [20, 79]. The threshold for this bias to become negligible is dependent on both the life history of the plant and how many plants were sampled. Ramula and coauthors determined that for the monocarpic perennial Cirsium palustre both the matrix model and the integral projection model are unbiased for a data set consisting of more than 300 records, resulting in nearly identical predicted values for population growth rate [79]. This claim is supported by my result that the matrix model and IPM do not produce significantly different values for population growth rate in dense stands with numerous individuals ( $n=554$ plants). However, due to the large size of the confidence intervals,
my analysis does not support the conclusion made by Ramula and coauthors that bias in estimated population growth rate becomes significant for smaller data sets.

When parameterized using a small demographic data set ( $n=103$ plants), the bootstrap distributions derived from each model have large variance [79]. The large variance in each estimate for population growth rate makes it difficult to derive strong conclusions from either model. Ramula and coauthors reported that, for small demographic data sets (fewer than 300 plants), IPMs have lower variance than matrix models in estimated population growth rate [79]. However, I did not observe this result for the open stand data set. For both the matrix model and the IPM the bootstrap confidence interval for dense stands is entirely within the bootstrap confidence interval for open stands.

I conclude that for both open and dense stands the management priorities recommended by Hüls and coauthors are unaffected by their decision to use a matrix model [39]. Specifically, the population growth rates predicted by the IPM support the authors' observation that open stands do not have a higher intrinsic population growth rate than dense stands.

The results of this chapter support the conclusion made by Hüls and coauthors that the populations growth rate in open stands is not significantly higher than population growth rate in dense stands [39]. The lack of a significant difference between open and dense stands can be contrasted with barbed goatgrass (Aegilops triuncialis), a grass invasive in North America that has been shown to have similar growth rates between 'core' and 'edge' populations [92]. Unlike giant hogweed, the 'core' and 'edge' populations of barbed goatgrass are caused by soil conditions rather than management. Due to the rocky soil conditions, spikes of barbed goatgrass in edge populations have an unusual ability to move long distances and colonize new habitats. In contrast to giant hogweed, edge populations of barbed goatgrass represent the invasion front and should be prioritized for control.

More similar to giant hogweed, some invasive species have two distinct subpopulations caused by control measures. For example, invasive lionfish species (Pterois volitans and Pterois miles) in the Caribbean are controlled by culling; however, due to diving limitations only reefs up to 30 m deep may be targeted for control [5]. This inability to cull populations in mesophotic reefs creates two subpopulations, one with low population density and one with high populations density, similar to open and dense stands of giant hogweed. Andradi-Brown and coauthors found that the ability for lionfish to quickly repopulate shallow reefs post-culling undermined management efforts. The authors concluded that enhancing management efforts to include culling in mesophotic reefs is necessary to control the species. Although giant hogweed is sessile and thus will replenish managed sites more slowly, expanding management to dense stands may be necessary to control the species.

Even though my results support the conclusions of Hüls and coauthors [39], they dispute the applicability of the findings of Ramula and coauthors towards the development of management regimes for invasive or endangered populations [79]. Ramula and coauthors concluded that small demographic data sets, such as the open stand data set, will have a biased population growth rate which may in turn affect management decisions. However, I found that the small size of the open stand data set makes it difficult to find significant differences in population growth rate. These results suggest that the decision to use an integral projection model may not reduce the difficulty in drawing conclusions from small demographic data sets in general.

The strength of the integral projection model approach is the use of continuous vital rate functions to construct the kernel. Unfortunately, this regression-based approach to describe population dynamics is dependent on the how well the vital rate functions approximate reality. I assumed that linear regression with normally distributed residuals would adequately describe plant growth. I tested these assumptions in section 2.2.4 and found that the fit could be improved upon by adding nonlinearities or by allowing residual errors to follow a different distribution.

Dahlgren and coauthors demonstrated that nonlinearities in the fitted growth function could have a significant impact on predicted population growth rate [19]. To address this concern I parameterized the IPM with each nonlinear model that had a lower AIC value than the linear model as well as a generalized additive model [82]. The population growth rates predicted by each of these models did not differ significantly from the population growth rate predicted by the IPM fit with linear growth. Similarly, the linear model with non-normal error distribution and generalized additive model with non-normal error distribution did not yield different predictions for population growth rate than the linear model used throughout this chapter.

Although predicted population growth rate was robust to changes in the growth fit and the error distribution, the selection of the linear fit with normal errors could affect other results of this chapter. The stable size distributions and the bootstrap confidence intervals for predicted population growth rate are both dependent on the form used for growth. The use of a nonlinear function to describe growth or relaxing the assumption that errors are normally distributed may affect the conclusions in this chapter; however, further investigation of the impact of growth fit is beyond the scope of this chapter.

In addition to the potential issue with growth, the fit for recruitment may be problematic. Since only eight recruits were present in open stands in 2003 [39], there is structural uncertainty in the recruit size distribution which may bias the estimate of population growth rate or stable size distribution. I used a log-normal distribution for both data sets
since it was a good fit for recruit size distribution in dense stands and since it is a commonly used distribution to describe recruit size [55]. However, the recruit size distribution may have a different form in open stands due to the difference in competitive pressure.

The uncertainty in recruit size is not the only issue with recruitment. In addition to the low number of recruits and uncertainty in recruit size distribution, there is uncertainty in how the number of flowering plants relates to the expected number of recruits. In particular, there were a few plots in the Hüls data set in which recruits were recorded in 2003 yet no flowering plants were recorded in 2002. These seeds may have originated from flowering plants just outside the plot or may have come from the seed bank [51, 57]. In either case, due to the unknown origin of these recruits we cannot expect the number of flowering plants to strongly correlate with the number of recruits in each plot.

As in the matrix model, the IPM predicts that populations are in decline. This decline was explained by Hüls and coauthors as being caused by drought conditions [39]. Hüls and coauthors found that populations rebounded in the 2004 growing season under normal weather conditions with the matrix model predicting population growth rates well above 1. Unfortunately, it is not possible to find the corresponding population growth rate using the IPM since plants were classified on site during the 2004 growing season. This dependence of population growth rate on realized weather conditions highlights the importance of multiyear studies of invasive populations since without the 2004 census Hüls and coauthors may have underestimated the risk of invasion by giant hogweed.

The low predicted population growth rates for giant hogweed due to the 2003 drought is not unusual; however, it is also common for invasive species to benefit from unusual weather conditions. Jackson reports that invasion of California grasslands by annual Mediterranean grasses was facilitated by drought conditions [41]. The author's hypothesis was later supported by a mechanistic resource availability model developed by Evarard and coauthors that finds that non-native annual grass species, such as great brome (bromus diandrus), competitively exclude perennial grasses native to Californian grasslands under drought conditions [31]. Subsequent research by Kimball and coauthors found that drought slowed recovery of Californian shrubland during postfire succession which may allow for the conversion of shrubland to grassland by non-native annual grasses [46].

The use of an integral projection model to estimate population growth rate for giant hogweed populations in this chapter indicates that open stands do not have a larger growth rate than dense stands, supporting the conclusions of Hüls and coauthors [39] that open stands are suppressed by biotic factors such as mowing or grazing. Furthermore, I found a practical limitation to the results of Ramula and coauthors [79] since the variance in bootstrap estimates for population growth rate is correlated with the magnitude of bias in
population growth rate. The high variance makes it difficult for significant differences to be reported for small demographic data sets. This limitation is exacerbated by the difficulty in computing bootstrap statistics using an IPM with only a few recruits.

## Chapter 3

## The Impact of Density-Dependent Flowering on H. mantegazzianum Seed Production

### 3.1 Introduction

The growth of a plant population must necessarily occur in an environment with finite resources. As a result of this resource limitation, the vital rate functions that govern plant population dynamics will vary as population density changes. Proper understanding of the transient and long-term dynamics of a population must therefore account for the impact of intraspecific competition on plant survival, growth, reproduction, fecundity, germination, and establishment. Although competition is important to account for, not each of these vital rates will have a significant impact on spread and their inclusion may needlessly complicate the model.

In the previous chapter I used a density-independent integral projection model to compare giant hogweed populations that have been classified as either open or dense stands. This classification was based on the level of interspecific competition giant hogweed plants face [39]. However, it has been reported by Pergl and coauthors that flowering is delayed in the presence of conspecifics [73] and it has been reported by Hüls and coauthors that competition from adult conspecifics results in longer generation times and a different size distribution [39]. In this chapter I extend the integral projection model to account for intraspecific competition. Before I describe how this is done, I must review previously published density-dependent models for giant hogweed.

Previous models for giant hogweed population dynamics have either ignored intraspecific density-dependence [39, 73], assumed there is a ceiling capacity below which populations are unaffected [63], or used a ceiling capacity with a smoothing function for survival [56]. However, the imposition of a ceiling capacity is not justified and will affect population dynamics. In particular, the predicted number of seeds produced will differ if a different form of density-dependence is used in the model. This difference in predicted seed set may have implications on the rate of spread of the species [68].

The most common form of density-dependence in integral projection models is recruitment limitation [81]. Recruitment limitation is most commonly modelled since small plants are more vulnerable than adult plants; however, another factor contributing to the focus on recruitment limitation is the relative ease of analysing models with only recruitment limitation when compared to models that account for adult density [81, 22, 23, 21]. The dearth of IPMs that allow for competition among adult conspecifics provides additional motivation for this chapter. This chapter will therefore have two models, one that only accounts for competition among recruits and a second that also includes adult competition.

Each recruit will compete for available microsites and resources in order to establish. I incorporate recruitment limitation in each of the models presented in this chapter. I then extend one of the models to allow for adult plants to compete with one another. This competition may result in reduced survival, growth, or probability of reproduction as adult plants compete for resources [98]. It will be determined during parameterization which density-dependent effects must be accounted for. However, if any of the aforementioned vital rates are affected by density then the total seed set may be affected.

It has previously been shown that the total seed set has an impact on the rate of spread of species, particularly when habitats are fragmented [68]. The rate of spread for giant hogweed is highly variable and sensitive to long-distance dispersal events that colonize disturbed sites [59, 72, 56]. The rate of spread is also dependent largely on the dispersal of seeds along corridors such as roads and waterways [93]. The accurate prediction of seed production in giant hogweed populations is therefore essential for realistic predictions of the spread of the species.

However, Pergl and coauthors [73] reported that flowering of giant hogweed plants is delayed in stands with high population density, which will in turn affect the seed production. Therefore, I expect that a negative relation between population density and the probability of flowering may initially result in higher population growth rates if there are significantly more seeds produced immediately following invasion as compared to when the populations have reached carrying capacity. I also expect that population growth rates over time will be lower if there are fewer seeds once populations have reached carrying capacity.

In this chapter I test this hypothesis using a structured population model of giant hogweed that includes negative density-dependence in two different forms; the first form is competition between seedlings for available microsites while the second form is the negative relation between the probability of flowering and population density. I compare the seed number predicted by the model with and without the delay in flowering and find that significantly fewer seeds are produced over time when the delay in flowering is modelled. Additionally, I find the equilibria for each model and the dominant eigenvalues for their Jacobians in order to determine if the course of invasion differs qualitatively. These differences in predicted seed production may in turn predict a different rate of spread for the species when competition among adult plants is accounted for.

### 3.2 Methods

### 3.2.1 Study species

Giant hogweed is a monocarpic perennial herb which may grow up to 5.5 m tall [69]. The species is native to the Caucasus and has successfully invaded much of Europe and North America [69]. The invasiveness of giant hogweed is bolstered by the extremely large seed set consisting of thousands or tens of thousands of seeds which may disperse along roads and waterways [56]. Adult plants outshade competitors and may potentially produce allelopathic substances to encourage the establishment of their offspring [100]; however, the large seed set left behind leads to high intraspecific competition among seedlings. Populations are typically not seed limited, but in patchy environments seed limitation has the potential to impact invasion speed [68].

### 3.2.2 Experimental data collection

I used the annual transition data collected in Germany between 2002 and 2003 by Dr. Jörg Hüls and coauthors [39] to parameterize adult population dynamics (Table 2.1). Hüls and coauthors collected measurements of height, the petiole diameter for the largest leaf, and laminar width of the largest leaf for vegetative plants in each year. As was done in the previous chapter, I used height as the measure of plant size.

Mean seed production was also recorded in the original study; however, seed production differed greatly from most previously published estimates [66, 74, 77, 94], yet agreed with estimates from giant hogweed populations in Ireland [10, 75]. However, it is speculated
that the estimates of seed production from the study in Ireland are overestimated due to the author estimating fruit number by counting both male and female flowers [75]. Rather than use the estimates produced by Hüls and coauthors I estimated mean seed production using a meta-analysis from the literature. Mean seed production was estimated from the weighted mean of estimates from four sources from the literature (Table 3.1, excluding Caffrey).

Monthly seedling survival data were collected in the original study by Hüls and coauthors [39]. Plots with an area of $0.1 \mathrm{~m}^{2}$ were saturated with seeds so that further addition of seeds would not yield more seedlings (Jörg Hüls, personal communication). Hüls and coauthors recorded the number of surviving seedlings each month. I used the average number of seedlings measured in July as the maximum number of seedlings per plot. However, since plots were saturated with seeds I needed additional data to parameterize a recruitment function with few seeds.

To finish parametrizing the model, Grguric and Cuddington set up 3 sites in Ontario with 4 quadrants each to estimate germination (Table 3.2). Fifty seeds were sowed in each plot in Fall of 2016 and the number of resulting seedlings recorded in Spring of 2017.

Table 3.1: Estimates of seed production from the literature.

| Author(s) | Location | Mean seeds | N | Source |
| :--- | :--- | ---: | :---: | :---: |
| Caffrey | Ireland | 41202 | 80 | $[10]$ |
| Ochsmann | Germany | 9695 | 33 | $[66]$ |
| Perglova | Czech Republic | 20671 | 98 | $[74]$ |
| Pysek | Czech Republic | 16140 | 8 | $[77]$ |
| Tiley, Philp | Scotland | 15729 | 4 | $[94]$ |
| Weighted mean (excluding Caffrey) |  | 17746 |  |  |

### 3.2.3 Towards a density-dependent integral projection model

## State variables

In order to account for density-dependence I first needed mathematical definitions for density. For competition between recruits I used the total number of seeds at time t, $n_{s}(t)$, while for adult competition I used the total biomass in the plot, $b(t)$. The decision to use total plot biomass as a measure of competition was based on previous studies

Table 3.2: Seedling germination data collected in Ontario.

| Location | 2016 census | 2017 census | Quadrant | 2016 seeds | 2017 seedlings |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Airport Rd | Oct 18 | Apr 28 | 1 | 50 | 33 |
|  |  |  | 2 | 50 | 28 |
|  |  |  | 3 | 50 | 26 |
|  |  |  | 4 | 50 | 23 |
| Hwy 9 | Oct 18 | Apr 28 | 1 | 50 | 10 |
|  |  |  | 2 | 50 | 21 |
|  |  |  | 3 | 50 | 22 |
|  |  |  | 4 | 50 | 12 |
| Woodstock | Oct 19 | May 12 | 1 | 50 | 4 |
|  |  |  | 2 | 50 | 5 |
|  |  |  | 3 | 50 | 6 |
|  |  |  | 4 | 50 | 3 |

that show that the relative growth rates of target plants are significantly affected by the biomass of neighbouring plants [32, 33, 34, 80]. Since flowering plants will be competing for resources I needed to track the size distribution of both vegetative and flowering plants to determine total biomass. I define $n_{v}(x, t)$ to be the size distribution of vegetative plants at time $t, n_{r}(x, t)$ to be the size distribution of reproductive plants at time $t$, and $n(x, t)=\left[n_{v}(x, t), n_{r}(x, t)\right]$ to be the total state of the population at time $t$.

I assumed that the number of seeds produced is proportional to the number of flowering plants. Previous authors have observed that the number of flowers will increase with plant size; however, the majority of the excess flowers do not produce fruit [75]. Let $r_{t}(t)$ be the total number of flowering plants at time $t$. Then $r_{t}(t)$ may be found by integrating over the size distribution of flowering plants, $r_{t}(t)=\int_{L}^{U} n_{r}(x, t) \mathrm{d} x$, where $L$ and $U$ are the lower and upper bounds for individual size, respectively. The total number of seeds at time $t$ may then be expressed as an integral, $n_{s}(t)=m_{s} r_{t}(t)=\int_{L}^{U} m_{s} n_{r}(x, t) \mathrm{d} x$, where $m_{s}$ is the mean number of seeds a reproductive plant produces (Table 3.1). This mean number of seeds produced by each plant is assumed to be independent of plant density [73] and independent of plant size [75].

The measure of adult competition used in the model was necessarily a function of plant heights within each plot. This restriction was due to limitations in the data set collected by Hüls and coauthors [39] and due to height being the only state variable measured in the
model. I assumed that competition between adults is size-symmetric and dependent on the total biomass of the population [87, 97, 98]. Unfortunately, measurements of biomass were not collected by Hüls and coauthors. Instead, the total biomass of adult plants in the population was estimated using the allometric scaling law of West, Brown, and Enquist [28, 29, 30, 84, 99]. This law states that the biomass of a terrestrial plant, $b_{0}$, scales with its height, $x_{0}$, to the fourth power. I applied this relation to vegetative and reproductive plants at each height and integrated in order to find the total biomass at time $t$ :

$$
\begin{equation*}
b(t)=\int_{L}^{U} x^{4}\left[n_{v}(x, t)+n_{r}(x, t)\right] \mathrm{d} x . \tag{3.1}
\end{equation*}
$$

The scaling law of West, Brown, and Enquist has previously been criticized for making assumptions that do not necessarily hold for every plant species [49, 50]. Instead, other authors give a range between 3 and 6 for the exponent [3]. I reparameterized the model using the values of 3 and 6 for the scaling exponent and reran the simulations in the results section to determine if this uncertainty in the allometric scaling law affects my results. These findings are briefly mentioned in the discussion.

## Vital rate functions

Excluding immigration, there are three processes that describe giant hogweed population dynamics: the survival and growth of vegetative plants, the transition of vegetative plants to their reproductive state, and the production of recruits by the reproductive plants. These processes of survival-growth, reproduction, and fecundity may be governed by integral kernels similar to the density-independent integral projection model from chapter 2.

In the recruit-limited model I assumed that the number of recruits is a monotonically increasing, bounded function of the number of number of seeds produced in the plot. The fecundity kernel, $F$, is therefore a function of the number of seeds, $n_{s}$. Specifically, the function that gives the number of surviving recruits has a Michaelis-Menten form [22]. The survival-growth and reproduction kernels are assumed to be density-independent as in chapter 2. The recruit-limited model has the form:

$$
\begin{align*}
& n_{v}(y, t+1)=\int_{L}^{U} P(y, x) n_{v}(x, t) \mathrm{d} x+\int_{L}^{U} F\left(y, x, n_{s}\right) n_{r}(x, t) \mathrm{d} x  \tag{3.2}\\
& n_{r}(y, t+1)=\int_{L}^{U} R(y, x) n_{v}(x, t) \mathrm{d} x \tag{3.3}
\end{align*}
$$

where $P(y, x)$ is the survival-growth kernel, $R(y, x)$ is the reproduction kernel, and $F\left(y, x, n_{s}\right)$ is the fecundity kernel.

In the adult competition model I allowed the fecundity kernel to depend on the total number of seeds. However, this model differs from the recruit-limited model in that I allowed the functions describing survival, growth, and the probability of reproduction to depend on total plant biomass. The general form of the adult competition IPM is:

$$
\begin{align*}
& n_{v}(y, t+1)=\int_{L}^{U} P(y, x, b) n_{v}(x, t) \mathrm{d} x+\int_{L}^{U} F\left(y, x, n_{s}\right) n_{r}(x, t) \mathrm{d} x  \tag{3.4}\\
& n_{r}(y, t+1)=\int_{L}^{U} R(y, x, b) n_{v}(x, t) \mathrm{d} x \tag{3.5}
\end{align*}
$$

where $P(y, x, b)$ is the survival-growth kernel, $R(y, x, b)$ is the reproduction kernel, and $F\left(y, x, n_{s}\right)$ is the fecundity kernel.

## Determining the integral kernels

The survival-growth kernel for the adult competition model has the same form as survivalgrowth kernels in other IPMs for monocarpic perennials [14], albeit with vital rates dependent on adult plant biomass:

$$
\begin{equation*}
P(y, x, b)=p_{s}(x, b)\left[1-p_{r}(x, b)\right] G_{v}(y, x, b), \tag{3.6}
\end{equation*}
$$

where $p_{s}(x, b)$ is the probability a plant of size $x$ in a plot of biomass $b$ survives to the next census, $p_{r}(x, b)$ is the probability a plant of size $x$ in a plot of biomass $b$ reproduces before the next census, and $G_{v}(y, x, b)$ is the probability a vegetative plant of size $x$ in a plot of biomass $b$ that remains vegetative grows to size $y$ in the next census. The survivalgrowth kernel for the recruit-limited model is identical except for the lack of dependence
on biomass.
The reproduction kernel for the adult competition model is the probability a vegetative plant survives, reproduces, and grows to a given size:

$$
\begin{equation*}
R(y, x, b)=p_{s}(x, b) p_{r}(x, b) G_{r}(y, x, b), \tag{3.7}
\end{equation*}
$$

where $G_{r}(y, x, b)$ is the probability a plant of size $x$ in a plot of biomass $b$ that reproduces grows to size $y$ the next census. The reproduction kernel for the recruit-limited model is identical except for the lack of dependence on biomass.

Lastly, I found the fecundity kernel for both the recruit-limited model and the adult competition model. This kernel was found by first integrating over the fecundity kernel multiplied by the distribution of reproductive plants. This integral must be equal to the total number of recruits, $f_{e}\left(n_{s}\right)$, multiplied by the recruit size distribution, $c_{0}(y)$. Therefore $\int_{L}^{U} F\left(y, x, n_{s}(t)\right) n_{r}(x, t) \mathrm{d} x=g_{e}\left(n_{s}(t)\right) n_{s}(t) c_{0}(y)=\int_{L}^{U} m_{s} g_{e}\left(n_{s}(t)\right) c_{0}(y) n_{r}(x, t) \mathrm{d} x$, where $g_{e}\left(n_{s}\right)$ is the probability a seedling establishes when $n_{s}$ seeds are produced. This equality implies that:

$$
\begin{equation*}
F\left(y, x, n_{s}\right)=m_{s} g_{e}\left(n_{s}(t)\right) c_{0}(y) \tag{3.8}
\end{equation*}
$$

### 3.2.4 Parameterization

## Bounds on plant size

The lower bound, $L$, was chosen to be the same as the lower bound in the densityindependent IPM, $L=5 \mathrm{~cm}$. I selected the upper bound of $U=550 \mathrm{~cm}$ since that value is a commonly reported upper bound for flowering giant hogweed plants [78, 69].

## Probability of survival and reproduction

I assumed that the probability of survival and the probability of reproduction were both density-independent functions for the recruit-limited model. I used logistic regression with height as a covariate to fit these vital rate functions. However, this assumption is not necessarily valid since intraspecific competition for resources may affect plant survival or
the timing of reproduction. Therefore, for the adult competition model I fit the probability of survival and probability of reproduction using logistic regression with total plot biomass as a covariate in addition to plant height.

I compared the AIC values of the density-independent and density-dependent fits in order to determine if the vital rates in the adult competition model should depend on adult biomass. I found that the probability of survival had no improvement in fit when biomass was accounted for (Table 3.3). I therefore used the same density-independent function as was used in the recruit-limited model for the probability of survival. However, including biomass as a covariate in regression for the probability of reproduction did yield a significantly better fit than the density-independent fit $(\triangle \mathrm{AIC}=13, p=0.013$ from the likelihood ratio test). Therefore, for the adult competition model I used a function for the probability of reproduction that depends on the biomass of neighbouring plants (Table 3.4).

Table 3.3: Impact of biomass and site-specific effects on fits for vital rate functions for giant hogweed. Vital rates were each fit twice - the first fit does not use biomass as a covariate in regression while the second fit does use biomass as a covariate. The best performing fit for each data set is in bold. All data collected by Hüls and coauthors [39].

| Vital Rate | Density- <br> dependent? | df | AIC |
| :--- | :---: | :---: | :---: |
| Survival $\left(p_{s}\right)$ | No | 2 | $\mathbf{4 8 3}$ |
|  | Yes | 3 | $\mathbf{4 8 3}$ |
| Reproduction $\left(p_{r}\right)$ | No | 2 | 383 |
|  | Yes | 3 | $\mathbf{3 7 0}$ |
| Growth of vegetative plants $\left(G_{v}\right)$ | No | 2 | $\mathbf{2 7 3 9}$ |
|  | Yes | 3 | $\mathbf{2 7 3 9}$ |
| Growth of reproductive plants $\left(G_{r}\right)$ | No | 2 | 1117 |
|  | Yes | 3 | $\mathbf{1 1 1 5}$ |

Table 3.4: The final forms of the vital rate functions. All functions were fit using data collected by Hüls and coauthors [39]. RL=recruit-limited and AC=adult competition. Standard error is indicated in parentheses. The Michaelis constant for mean recruits was fit using data collected by Cuddington and Grguric.

| Function | Notes | Fitted function |
| :--- | :--- | :---: |
| Survival $\left(p_{s}\right)$ |  | $\operatorname{Logit}\left(p_{s}(x)\right)=-0.284_{(0.235)}+0.023_{(0.004)} x$ |
|  | RL | $\operatorname{Logit}\left(p_{r}(x)\right)=-3.340_{(0.349)}+0.028_{(0.004)} x$ |
| Reproduction $\left(p_{r}\right)$ | AC | $\operatorname{Logit}\left(p_{r}(x, b)\right)=-2.889_{(0.370)}+0.030_{(0.004)} x-0.003_{(0.0008)} b$ |
|  | V | $G_{v}(y, x)=36.68_{(4.96)}+0.66_{(0.06)} x+\mathcal{N}\left(0,37.52^{2}\right)$ |
| Growth $(G)$ | R | $G_{r}(y, x)=217.29_{(10.85)}+0.60_{(0.09)} x+\mathcal{N}\left(0,36.30^{2}\right)$ |
| Mean recruits $\left(f_{e}\right)$ |  | $f_{e}\left(n_{s}\right)=\frac{127.5}{n_{s}+273.4} n_{s}$ |
| Recruit size $\left(c_{0}\right)$ |  | $\ln (y) \sim \mathcal{N}\left(3.82,0.55^{2}\right)$ |



Figure 3.1: Data and best-fit functions for relationships between plant height and vital rates. Survival probability (a), expected height of vegetative (b, black dots) and reproductive plants (b, unfilled rhombi), and probabilities of reproduction for the recruit-limited model (c) and adult competition model with zero biomass (d). All data collected by Hüls and coauthors [39].

## Inter-annual growth

I fit seven common functional forms to describe the growth of vegetative and reproductive plants [70]. I then compared each fit using their AIC values to determine which form to use. I found that the discrete logistic and power law fits had a slightly lower AIC value than the linear model for growth of vegetative plants. However, this difference is very small $(\Delta \mathrm{AIC} \leq 3)$ and so I opted to use the linear model. For the growth of reproductive plants, I found that the linear fit performed better than each nonlinear fit. I therefore selected
linear functions to describe the growth rates of vegetative and reproductive plants.
To test the possibility that density may affect growth rates I fit functions for the growth of vegetative and reproductive plants using linear regression with plant height and population biomass as covariates. Although the AIC value for the function describing the growth of vegetative plants was not improved by using density as a covariate, the function describing the growth of reproductive plants did improve slightly (Table 3.3). However, since this improvement was small $(\Delta \mathrm{AIC}=2)$, I decided to use the linear model with only height as a covariate for both growth functions in both models.

In order to evaluate if the linear functions are suitable, I tested both fits for heteroskedasticity and normality of residuals. The Breusch-Pagan test indicated that the hypothesis that variance is constant cannot be rejected ( $p=0.0887$ for vegetative plant growth and $p=0.06978$ for flowering plant growth) and therefore the assumptions of homoskedasticity are not violated. Similarly, the Shapiro-Wilk test does not reject the hypothesis that residuals are normally distributed ( $p=0.1223$ for vegetative plant growth and 0.2917 for flowering plant growth). I therefore concluded that the growth kernel is a Gaussian conditional probability distribution with mean scaling linearly with height and constant variance.

## Expected number of recruits

Like Eager and coauthors [22], I used a Michaelis-Menten function to describe the expected number of recruits for a given number of seeds. The Michaelis-Menten function was derived from first principles by Eager and coauthors by assuming seeds compete for available microsites in which the seeds may establish. Let $n_{s}$ be the number of seeds, $f_{e}\left(n_{s}\right)$ be the total number of recruits that will result from $n_{s}$ seeds, $\alpha$ be the maximum number of seedlings a plot may sustain, $\beta$ be the number of seeds that would result in $\alpha / 2$ seedlings, and $g_{e}\left(n_{s}\right)$ be the probability a seed germinates and establishes given $n_{s}$ seeds. The form used for recruitment is:

$$
\begin{equation*}
f_{e}\left(n_{s}\right)=\frac{\alpha}{\beta+n_{s}} n_{s}=g_{e}\left(n_{s}\right) n_{s} \tag{3.9}
\end{equation*}
$$

I used a combination of data collected by Hüls and coauthors [39] and data collected by Cuddington and Grguric to parameterize the recruitment function.

Hüls and coauthors [39] conducted seed sowing experiments in which $0.1 \mathrm{~m}^{2}$ plots were saturated with giant hogweed seeds. The exact number of seeds used in the experiments
is unknown and so I cannot use regression to fit the Michaelis-Menten function. However, by assuming a sufficiently large number of seeds were used in each experiment, I can assume that the mean number of seedlings in each plot is the expected maximum number of seedlings that may establish in $0.1 \mathrm{~m}^{2}$. The mean number of seedlings that established in these plots was 5.1. Since the plots in which Hüls and coauthors collected all their other data each have an area of $2.5 \mathrm{~m}^{2}$, this translates to a maximum seedling capacity of $\alpha=127.5$.

In order to determine $\beta$, I used the data collected by Cuddington and Grguric in which 50 seeds were each sowed in twelve $1 \mathrm{~m}^{2}$ plots (Table 3.2). The mean number of seedlings which resulted were 16 seedlings $/ \mathrm{m}^{2}$. In a plot with an area of $2.5 \mathrm{~m}^{2}$, this is equivalent to 125 seeds sowed and a resulting 40 seedlings. Therefore:

$$
\begin{equation*}
f_{e}(125)=\frac{127.5}{\beta+125} 125=40 \tag{3.10}
\end{equation*}
$$

which implies that $\beta=273.4375$.

## Recruit size distribution

As was done in chapter 2, I assumed that recruit size follows a log-normal distribution. I tested this assumption by log-transforming recruit height and using the Shapiro-Wilk test for normality. The Shapiro-Wilk test did not reject the hypothesis that the log-transformed values are normally distributed ( $p=0.58$ ). I denote the recruit size distribution by $c_{0}(y)$ (Table 3.4). Unintentional eviction was prevented by truncating the distribution at $L=5$ cm and $U=550 \mathrm{~cm}$ and normalizing.


Figure 3.2: Number of seedlings as a function of the seed set in the plot (a) and histogram of observed distribution of recruit size along with the fitted log-normal distribution for stands of giant hogweed in the Hesse region of Germany ( $n=75, \mathrm{~b}$ ).

## Site-specific effects

I used mixed-effects modelling with a random intercept to test whether any vital rate functions were significantly affected by sampling from different sites [7]. The AIC values for the mixed-effects models for survival, reproduction, and the growth of vegetative plants were higher than their corresponding fixed-effects models. The AIC value for the mixed-effects model for the growth of reproductive plants was lower than the AIC value for the corresponding fixed-effects model. However, this difference was not significant as determined by the likelihood-ratio test ( $\Delta \mathrm{AIC}<2 ; p=0.07345$ ). Therefore, I did not use mixed-effects modelling in the final model.

## Outliers

I removed the same two outliers that were excluded from the density-independent model in the previous chapter. However, there was the potential for more outliers to be identified since the density-dependent model must necessarily track the heights of reproductive plants. I found one reproductive plant that had the magnitude of its residual for growth more than 4 times larger than the standard deviation of the residuals for the function describing the growth of reproductive plants. This plant appears to have had its height recorded incorrectly and so I excluded it from the density-dependent model.

### 3.2.5 Simulation results

The initial population in each simulation was one seed and no vegetative plants nor flowering plants. These initial conditions were selected to simulate the local population dynamics of a population from the start of an invasion. Populations were simulated for 30 years to determine how seed production changes over the course of an invasion.

The number of seeds produced in the second year were compared to determine how short distance dispersal would be affected. The number of seeds produced in years 3-7 were compared to determine if seed production was significantly affected early in the invasion. Finally, the number of seeds produced annually once populations became established were compared to estimate the relative likelihood of overcoming significant barriers to dispersal, such as forests or managed sites.

### 3.2.6 Equilibrium size distribution and stability

I determined the existence of an equilibrium for the recruit-limited model and its stabliity using the same technique presented by Rebarber and coauthors [81]. I then modified this technique to find the equilibrium for the adult competition model. Stability of the equilibrium for the adult competition model was determined by computing the dominant eigenvalue of the Jacobian using methods modified from Ellner and Rees [27].

### 3.3 Results

### 3.3.1 Stability results for each model

The recruit-limited model satisfies the conditions outlined by Rebarber and coauthors to guarantee the existence of a globally asymptotically stable equilibrium [81]. However, the model that allows for adult competition does not satisfy these conditions and so simulations were performed to determine stability. Simulation results indicate that populations are oscillatory with a period of two years (Figure 3.3). This behaviour was confirmed by modifying the results of Rebarber and coauthors [81] to find the equilibrium for the adult competition model numerically. The Jacobian of this equilibrium has a dominant eigenvalue of -1.04 , confirming that the population dynamics around the equilibrium are unstable and oscillatory.

### 3.3.2 Seed production predicted by each model

The expected number of seeds produced in the second year is predicted to be 1037 seeds in the recruit-limited model and 1504 seeds in the adult competition model. Similarly, the expected number of seeds in the third year is predicted to be 1013 seeds in the recruitlimited model and 1220 seeds in the adult competition model. A second generation of plants may sprout by the fourth year, bringing the expected number of seeds up to 226142 in the recruit-limited model and 313768 in the adult competition model. However, none of these increases in seed production are significant (Table 3.5). Furthermore, these increases in seed production are transient behaviour.


Figure 3.3: Number of seeds in each year predicted by the recruit-limited model (circles) and the model with adult competition (rhombi).

Starting in the fifth year, the density-dependent probability of flowering results in

Table 3.5: Bootstrap mean seed production and confidence intervals for years $2,3,4,5,6$, 7 , 29, and 30 . $\mathrm{RL}=$ recruit-limited and $\mathrm{AC}=$ adult competition. 500 bootstrap simulations were performed for each model.

| Year | Model | Mean seeds | Bootstrap CI | Plot biomass | Flowering Plants |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | RL | 1038 | $(745,1344)$ | 4 | 0.059 |
|  | AC | 1501 | $(1020,1981)$ | 5 | 0.086 |
| 3 | RL | 1012 | $(823,1215)$ | 53 | 0.057 |
|  | AC | 1218 | $(1036,1414)$ | 57 | 0.069 |
| 4 | RL | 226298 | (150294, 307064) | 899 | 12.84 |
|  | AC | 311849 | (213527, 406929) | 1183 | 18.01 |
| 5 | RL | 443927 | (331568, 566017) | 1822 | 25.21 |
|  | AC | 63781 | (5744, 208389) | 426 | 2.50 |
| 6 | RL | 651931 | (507199, 802756) | 2717 | 36.99 |
|  | AC | 505463 | (291989, 767949) | 2524 | 29.85 |
| 7 | RL | 794209 | $(635208,963349)$ | 3352 | 45.03 |
|  | AC | 29916 | $(4,217112)$ | 423 | 0.14 |
| 29 | RL | 969529 | $(808318,1149654)$ | 4168 | 54.87 |
|  | AC | 2974 | (0.1384728, 4120) | 303 | 0.0013 |
| 30 | RL | 969529 | (808318, 1149654) | 4168 | 54.87 |
|  | AC | 832050 | (53407, 1261324) | 3930 | 46.44 |

oscillations in seed production in the adult competition model (Figure 3.3). In odd years, the seed production predicted by the model with adult competition is significantly lower than the seed production predicted by the recruit-limited model. However, in even years the number of seeds produced does not significantly differ between the models.

### 3.3.3 Differences in mean plant size

The mean height of vegetative plants was observed to be 75 cm with a $95 \%$ confidence interval of ( $71 \mathrm{~cm}, 80 \mathrm{~cm}$ ) in 2003. This confidence interval overlapped with the bootstrap equilibrium mean height of vegetative plants predicted by the recruit-limited model, as well as the mean plant heights predicted for each year in the adult competition model (Table 3.6). These bootstrap confidence intervals indicate that there is no significant difference between the observed mean height and the mean heights predicted by either
model.
The mean height of reproductive plants was observed to be 283 cm with a $95 \%$ confidence interval of ( $275 \mathrm{~cm}, 291 \mathrm{~cm}$ ) in 2003 (Table 3.6). This confidence interval overlapped with the bootstrap equilibrium mean height of reproductive plants predicted by the recruit-limited model. The confidence interval also overlapped with the corresponding $95 \%$ confidence interval predicted by the model with competition among adults. However, the bootstrap confidence interval for the mean height of reproductive plants predicted by the adult competition model in odd years did not overlap with the confidence interval of observed heights of reproductive plants in 2003. This result indicates there is a significant difference between the mean height of reproductive plants in odd years predicted by the model with adult competition when compared to the observed mean height of reproductive plants.

Table 3.6: Bootstrap observed mean height and predicted mean height of vegetative and reproductive plants. 500 bootstrap simulations were performed for each model. All heights presented in centimeters.

| Plant type | Source | Year | Bootstrap mean | Bootstrap CI |
| :--- | :--- | :---: | :---: | :---: |
| Vegetative | Observed | 2003 | 75 | $(71,80)$ |
|  | Recruit-limited | 30 | 69 | $(63,74)$ |
|  | Adult Competition | 29 | 76 | $(66,97)$ |
|  | Adult Competition | 30 | 83 | $(73,102)$ |
|  | Observed | 2003 | 283 | $(275,291)$ |
| Reproductive | Recruit-limited | 30 | 280 | $(271,290)$ |
|  | Adult Competition | 29 | 323 | $(295,389)$ |
|  | Adult Competition | 30 | 293 | $(279,327)$ |

### 3.4 Discussion

In this chapter I demonstrated that competition among adult giant hogweed plants results in significantly fewer seeds produced long-term, without a significant difference in initial seed production. I therefore suggest that modelling the delay in flowering may result in significantly lower rates of spread of giant hogweed [68]. In addition, the model with adult competition predicts oscillations in giant hogweed seed production. This prediction is in contrast with the model without adult competition which found that populations converge
to a stable equilibrium [81]. The number of seeds produced in the more fruitful years of the adult competition model is not significantly less than the number of seeds produced in the recruit-limited model. However, the number of seeds produced in less fruitful years of the adult competition model is significantly lower than the number of seeds predicted by the recruit-limited model. Therefore, the results of this chapter suggest the long-term population growth rate and rate of spread will be lower in the model with adult competition.

In order to parameterize this model, I had to combine data from a few different sources. I used data collected by Hüls and coauthors [39] to parameterize functions that describe survival, reproduction, and growth of adult plants. However, the open stand data set and dense stand data set were insufficient on their own to parameterize the model. Therefore, I combined the open and dense stand data sets to use in this chapter. Hüls and coauthors state that the stand types differ in the level of interspecific competition for each stand; however, in practice the authors classified stands based on the ground cover of giant hogweed. The differences between open and dense stands should therefore be accounted for by the measure of intraspecific competition used in this chapter, which will correlate with giant hogweed ground cover. Unfortunately, even after aggregating the data collected by Hüls and coauthors the data was still insufficient to parameterize the model.

To complete the data collection, Cuddington and Grguric collected recruit germination and survival data at three sites in Southern Ontario. The combination of this data set with the data set collected in Germany was necessary to parameterize the function for the expected number of recruits. Climatic conditions are somewhat different between the two locations and it is not known how results will change if recruitment data were collected at the same time as the adult plant demographic data. Although this model may not necessarily describe population dynamics in Germany or Canada, it will provide insights into the general course of giant hogweed establishment and long-term population dynamics. Unfortunately, the collection of adult data and recruit data in two separate locations does introduce an additional issue in the model.

As a result of the collection of demographic data in two locations, it was not possible to determine the effect that adult plants have on the survival and growth of recruits. This interaction is typically ignored in density-dependent models similar to the model presented in this chapter; however, competition between adult plants and seedlings will affect the equilibrium and may affect the results of my stability analysis [21]. In addition to the lack of adult-recruit interaction, I was unable to model the seed bank for this species [51, 57, 23]. The seed bank may help dampen or eliminate oscillations in the model if enough seeds persist in the seed bank to fill in after less fruitful years.

Finally, in order to get a measure of density I assumed that competition was propor-
tional to the biomass of all plants in the plot. This assumption is based off the work of West, Brown, and Enquist who derived allometric scaling relations between different measures of plant size [28, 29, 30, 84, 99]. Numerous researchers have criticised the result [49, 50, 3], arguing that the exponent will be species-dependent. Instead, the exponent for the model is likely in the range from 3 to $6[40,3]$.

I reran the simulations from subsection 3.3.2 with an exponent of 3 and with an exponent of 6 to determine if the uncertainty in the scaling law affects model results. Simulating population growth using the relation for biomass $M \propto h^{6}$ yields dynamics that are qualitatively the same as the results given in subsection 3.3.2; however, using the relation $M \propto h^{3}$ predicts that populations will initially reach a high density before settling into a low equilibrium density. Therefore, the uncertainty in the allometric scaling law could lead to drastically different conclusions. In addition, another assumption in my model may be causing oscillations.

I assumed that competition was symmetric since asymmetric competition is more difficult to model and more computationally expensive. However, it has previously been shown that models with symmetric competition may predict oscillations where similar models with asymmetric competition do not [67, 11]. To my knowledge oscillations have not been observed in giant hogweed populations, and so it is possible that the assumption that competition is symmetric may be leading to inaccurate population dynamics. However, even if the oscillations are an artefact of the choice of model, the impact of the density-dependent flowering will likely affect the conclusions of previously published matrix models and integral projection models.

Among matrix models, Pergl and coauthors developed a stochastic matrix model with a given probability for seeds to disperse a long distance away from the mother plant [72]. The authors then simulated the spread of giant hogweed with several different probabilities of long-distance dispersal and found that the values that best agreed with observed rates of spread is between $0.1 \%$ and $7.5 \%$. However, the matrices used in the simulation model were density-independent and did not account for the decrease in seed production that may occur in high density populations. This decrease in seed production may lead to different estimates of the probability of long-distance dispersal. Furthermore, Moenickes and Thiele developed a spatial matrix model to determine the factors that will have a significant impact on the rate of spread of the species [56]. However, the hypotheses the authors test each relate to various methods of spread, recruitment limitation, or succession, without concern for the importance of seed production on the rate of spread.

Among individual-based models, Wadsworth and Collingham produced an IBM that does not account for a decrease in seed production [96] and Nehrbass and coauthors produce
four individual-based models in their investigation of giant hogweed spread and control that assume a simple ceiling carrying capacity for the number of plants [63, 60, 61, 62]. In each of these models, population density may have a significant impact on the timing of flowering which in turn may have a significant impact on the results of these studies. However, further research is necessary to understand the implications my results may have on previous studies, if at all.

In each of these previous models, accounting for the delay in flowering may result in significantly fewer seeds being produced in some years. This decrease in seed production is likely to affect predictions of population spread. Pachepsky and Levine [68] observed that decreased seed production resulted in a lower rate of invasion in patchy habitats. This result is applicable to giant hogweed spread since the pattern of spread is typically patchy followed by infilling [62,54]. The density-dependent decrease in seed production could have an impact on previous models of giant hogweed spread.

Integrodifference models are commonly used to determine the rate of spread for invasive plants [48]. These models make use of a dispersal kernel to describe the dispersal of seeds and the dispersal kernel is commonly assumed to have exponentially bounded tails since leptokurtic dispersal kernels can lead to an unbounded rate of spread. With such dispersal kernels the invasion front moves at a constant speed. Furthermore, van den Bosch and coauthors demonstrate that in the absence of an Allee effect and the absence of longdistance dispersal the rate of spread is only dependent on the population growth rate at the front of the invasion [95]. However, most invasive plants have some long-distance dispersal that creates irregular invasion fronts and allow for plants that aren't at the front of the invasion to influence the rate of spread.

One particularly well-studied example of long-distance dispersal is the post-glacial migration of trees that occurred in the early Holocene. Clark used an integrodifference equation model to predict the rate of spread for various tree species and found that a leptokurtic dispersal kernel is necessary to explain the observed rates of spread [15]. More recently, Neubert and Caswell predicted the rate of spread of teasel in North America using a stagestructured integrodifference model, yet the predicted rate of $0.5639 \mathrm{~m} / \mathrm{yr}$ is well below the observed rate of $27 \mathrm{~km} / \mathrm{yr}$ [64]. The authors concluded that the large discrepancy between observed and predicted invasion speed was due to long-distance dispersal along waterways or multiple introductions of the species.

Similarly, Jongejans and coauthors used a stage-structured integrodifference model to predict the spread of musk thistle (Carduus nutans) in various non-native ranges [44]. The authors predicted spread rates of $7 \mathrm{~m} / \mathrm{yr}, 9 \mathrm{~m} / \mathrm{yr}$, and $29 \mathrm{~m} / \mathrm{yr}$ in Australia, Kansas, and New Zealand respectively. Jongejans and coauthors followed up with an integrodifference
model coupled with an IPM for musk thistle that predicted a rate of spread of $13.9 \mathrm{~m} / \mathrm{yr}$ in New Zealand [43]. However, each estimate differed greatly from the observed spread rates of $146 \mathrm{~m} / \mathrm{yr}$ in Pennsylvania. The authors concluded that the rate of spread in both of their models was underestimated since they did not account for long-distance dispersal. Rather than a uniformly advancing invasion front as predicted by Kot and coauthors [48], the front of musk thistle invasion is highly irregular due to long-distance dispersal. This pattern of invasion is relevant to this chapter since giant hogweed invasion may be driven by long-distance dispersal [72]. The density-dependent dynamics described in this chapter will be necessary to understand for giant hogweed spread due to it's sensitivity to long-distance dispersal.

The work in this chapter advances our understanding of the growth of giant hogweed in environments with limited resources. In particular, this work demonstrates that the predicted number of seeds produced by giant hogweed populations may be significantly affected if flowering is allowed to depend on density. This density-dependent flowering may in turn affect the rate of spread of the species and may need to be accounted for in future models of giant hogweed spread.

## Chapter 4

## Conclusions and Future Work

### 4.1 Conclusions

In this thesis I developed several models for giant hogweed population dynamics and investigated how model complexity may affect predicted population parameters.

From the results of chapter 2 I concluded that using an integral projection model rather than a matrix model did not change the main conclusion made by Hüls and coauthors [39]. That is, the integral projection model did not predict a significant difference in population growth rates between open and dense stands.

I concluded from the results of chapter 3 that allowing adult plants to compete with one another may result in significantly fewer seeds produced every other year. This decrease in seed production may lead to lower predictions for the rate of spread of the species [68]. Furthermore, as a result of the oscillations predicted by the adult competition model, the species may be forced to take advantage of invasion windows in order to establish new stands [37]. This result may impact previously published models of giant hogweed spread [96, 72, 56].

### 4.2 Future work

Endless research topics may be found by investigating how model complexity affects predicted population dynamics. A natural extension to the work presented in chapter 2 is to determine if population parameters other than population growth rate are affected by
the use of a matrix model or an integral projection model. Specifically, the predicted net reproductive ratio and predicted generation time for a population could be biased if predicted by a matrix model. Such results could further the research performed by Ramula and coauthors that proved that population growth rate may be biased by the discretization used to create some matrix models [79].

The work in chapter 3 was focussed on the impact that competition may have on total seed production and the possibility that the difference in seed production may lead to a higher or lower rate of spread. This work may be extended by developing stochastic, spatial integral projection models with and without adult competition and verifying if the rate of spread differs between the models. Such work would drastically improve the conclusions of chapter 3 by determining if a significantly different rate of spread is found, rather than simply speculating.

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## Appendices

## Appendix A

## Chapter 2 Code

## A. 1 Determining Population Growth Rates

```
```


### 

```
```


### 

## - Initialization

## - Initialization

# Clear workspace

# Clear workspace

rm(list=ls())
rm(list=ls())

# Library

# Library

library(MASS) \# Needed for 'fitdistr' function
library(MASS) \# Needed for 'fitdistr' function
library(lmtest) \# Needed for 'bptest' function
library(lmtest) \# Needed for 'bptest' function
library(mgcv) \# Needed for GAM fitting
library(mgcv) \# Needed for GAM fitting

# These are decisions that must be made

# These are decisions that must be made

exclude_outliers <- T
exclude_outliers <- T
growth_model <- "linear" \# can be "linear", "ricker", "skellam",
growth_model <- "linear" \# can be "linear", "ricker", "skellam",
"bevholt, "power", or "gam"
"bevholt, "power", or "gam"
normal_variance <- T
normal_variance <- T

### 

### 

## - Read in the data and define state variable

```
```


## - Read in the data and define state variable

```
```

20


```
# Data collected by Huls and coauthors:
# joerghuels@web.de
#
# Each row contains data for one plant
#
# The first column has an ID for the plant
# The first letter of the ID indicates the site: A=Allendorf,
# D=Druseltal, F=Frankenberg, K=Kassel, V=Viermunden
# The second letter indicates the stand type
# A,B,C,D,E=Dense stand; X=Open stand
# The Roman numerals indicate the plot
# The number indicates the plant
#
# Stage, height, leaf stem diameter, and leaf blade width
# were recorded for 2002 and 2003
# All measurements are in cm
# Stage 5 indicates the plant died that year
```



```
hmdata <- read.csv("./data/IPMdata.csv", na.strings=" "")
# I use height as the state variable
hmdata["h"] <- hmdata[["Height2002"]]
hmdata["h1"] <- hmdata [["Height2003"]]
###
## -- Remove Allendorf data, divide plots based on density
# The site is the first letter of the ID
hmdata["site"] <- substr(hmdata$ID,1,1) # first letter of
    individual ID
# Remove Allendorf data due to grazing damage
hmdata <- hmdata[hmdata$site!="A" ,]
# The density is the second letter of the ID
```

hmdata["dens"] <- ifelse(substr(hmdata\$ID, 2, 2) ="X", "Low", "High
\# Determine the plot number
hmdata["plot"] $<-$ substr (hmdata\$ID,1, as. numeric (lapply (as.
character (hmdata\$ID), nchar)) - 2 )
\#\#\#
\#\# - Survival, reproduction, and classification (i.e. seedling
or adult)
\# 1 means survived, 0 means did not survive, NA means flowered (
so of course it's dead)
hmdata["survived"] <- ifelse (hmdata [["Stage2002"]] $=4$, NA,
ifelse (hmdata [["Stage2003"]] =5, 0, 1))
\# 1 means reproduced, 0 means did not reproduce, NA means
individual died without reproducing
hmdata["reproduced"] <- ifelse (hmdata [["Stage2003"]] =4, 1,
ifelse (hmdata [["Stage2003"]] $=5$, NA, 0) )
\# If a plant was recorded in 2003 but not 2002 it was a seedling
in 2002
hmdata["class"] <- ifelse(is.na(hmdata[["Stage2002"]]), "Seedling
", "Adult")
\#\#\#
\#\# - Create another dataframe without the outlier
\# Help identify outliers
out1 $<-$ as.character (hmdata $[h m d a t a \$ h>80 \& ~ h m d a t a \$ h 1<10 \&!i s . n a($
hmdata\$h) \& ! is.na(hmdata\$h1),"ID"])
out $2<-$ as.character (hmdata $[h m d a t a \$ h<50 \& ~ h m d a t a \$ h 1>200 \& ~ h m d a t a \$$
reproduced $==0 \&!$ is.na(hmdata $\$ h) \&!$ is.na(hmdata $\$$ h1) ,"ID"])
out $3<-$ as.character (hmdata [is.na(hmdata $\$$ h) \& is.na(hmdata $\$$ h1),"
ID" ])

| 84 |  |
| :---: | :---: |
| 85 | \# Take a look at the outliers |
| 86 | hmdata [hmdata $\$$ ID $=$ out1, ] |
| 87 | hmdata [hmdata $\$$ ID $=$ out2, ] |
| 88 | hmdata[hmdata $\$$ ID $=$ out3, ] |
| 89 |  |
| 90 | \# Create a new data frame without them |
| 91 | hmdata_no_outlier <- hmdata[hmdata $\$$ ID!=out1 \& hmdata $\$$ ID!=out2 \& hmdata $\$$ ID! =out3, ] |
| 92 |  |
| 93 |  |
| 94 | \#\# |
| 95 | \#\# - Split single data frame into one for seedlings and one for |
| 96 |  |
| 97 | seedlingdata <- hmdata[hmdata ${ }^{\text {c class }}$ = "Seedling", ] |
| 98 | adultdata $<-$ hmdata[hmdata\$class!="Seedling", ] |
| 99 |  |
| 100 | ```seedlingdata_no_outlier <- hmdata_no_outlier[hmdata_no_outlier $ class="Seedling",]``` |
| 101 | adultdata_no_outlier $<-$ hmdata_no_outlier[hmdata_no_outlier $\$$ class!="Seedling" ,] |
| 102 |  |
| 103 |  |
| 104 | \#\#\# |
| 105 | \#\# - The data set to perform regression on |
| 106 |  |
| 107 | \# Which subset of data to use |
| 108 | if (exclude_outliers) \{ |
| 109 | dat $<-$ adultdata_no_outlier |
| 110 | $\}$ else $\{$ dat $<-$ adultdata $\}$ |
| 111 |  |
| 112 | dat_low <- dat [dat \$dens="Low", ] |
| 113 | dat_high <- dat[dat\$dens=" High", ] |
| 114 |  |
| 115 |  |
| 116 | \#\#\# |
| 117 | \#\# - Probability of survival |

146
\#\#\#
\#\# - Determine growth
\# I need to exclude flowering plants and dead plants
dat_adult <- dat[dat\$Stage2003!=4 \& dat\$Stage2003!=5,]
dat_adult_low <- dat_adult [dat_adult \$dens="Low", ]
dat_adult_high $<-$ dat_adult [dat_adult $\$$ dens=" High", $]$
\# Linear regression
if (growth_model_" linear") \{
grow_fit_low $<-\operatorname{lm}\left(h 1^{\sim} h\right.$, data=dat_adult_low)
grow_fit_high $<-\operatorname{lm}\left(h 1^{\sim} h\right.$, data=dat_adult_high $)$
\} else if (growth_model_"ricker") \{
rick_fun $<-$ function (x, coefs) $\{a<-\operatorname{coefs}[1] ; b<-\operatorname{coefs}[2]$;
return $(a * x * \exp (-b * x))\}$
grow_fit_low $<-$ nls (h1 $\sim_{\sim}^{r i c k}$ _fun $(h, c(a, b))$, data=dat_adult_low , start=list $(a=1, b=0.01))$
grow_fit_high <- nls (h1 ~ rick_fun(h, c(a,b)), data=dat_adult_ high, start $=$ list $(a=1, b=0.01))$
\} else if (growth_model="skellam") \{
skel_fun $<-$ function (x, coefs) $\{a<-\operatorname{coefs}[1] ; b<-\operatorname{coefs}[2] ;$ return $(a *(1-\exp (-b * x)))\}$
grow_fit_low $<-$ nls (h1 $\sim$ skel_fun $(h, c(a, b))$, data=dat_adult_low , start=list $(a=100, b=0.01))$
grow_fit_high <- nls (h1~ skel_fun (h, c (a, b) ), data=dat_adult_ high, start=list $(a=100, b=0.01))$
\} else if (growth_model="power") \{
powr_fun $<-$ function (x, coefs) $\{a<-\operatorname{coefs}[1] ; b<-\operatorname{coefs}[2] ;$ return $\left.\left(a * x^{\wedge} b\right)\right\}$
grow_fit_low $<-$ nls (h1 ~ powr_fun(h, c(a,b)), data=dat_adult_low , start=list $(a=1, b=1))$
grow_fit_high $<-$ nls (h1 ~ powr_fun(h, $c(a, b))$, data=dat_adult_ high, start=list $(a=1, b=1))$
\} else if (growth_model="bevholt") \{
bvht_fun $<-$ function (x, coefs) $\{a<-\operatorname{coefs}[1] ; b<-\operatorname{coefs}[2]$; return $(a * x /(b+x))\}$

```
    grow_fit_low <- nls(h1 ~ bvht_fun(h, c(a,b)), data=dat_adult_low
        , start=list(a=1,b=1))
    grow_fit_high <- nls(h1 ~ bvht_fun(h,c(a,b)), data=dat_adult_
        high, start=list (a=1,b=1))
} else if (growth_model="gam") {
    grow_fit_low <- gam(h1~s(h), data=dat_adult_low)
    grow_fit_high <- gam(h1~s(h), data=dat_adult_high)
} else {rm}
###
## -- Growth kernel
# Define growth kernel
g_k <- function(h1,h, dens) {
    # Define nice data frame to use
    newdata <- data.frame(h=h)
    # Two cases: open or dense stands
    if (dens="Low") {
        h1bar <- predict(grow_fit_low, newdata=newdata, type="
            response")
        res <- residuals(grow_fit_low)
        df_res <- df.residual(grow_fit_low)}
    else {
        h1bar <- predict(grow_fit_high, newdata=newdata, type="
                response")
        res <- residuals(grow_fit_high)
        df_res <- df.residual(grow_fit_high)}
    # I need standard error and estimated standard deviation
    sse <- sum(res^2)
    sdhat <- sqrt(sse/df_res)
    # Two cases: normal variance or nonparametric variance
    # see 'kernelExample.R' from Rees, Childs, Ellner (2014)
    if (normal_variance) {
```

```
            out <- dnorm(h1, mean=h1bar, sd=sdhat)
    } else {
            bw <- bw.SJ(res); alpha <- sdhat/sqrt(sdhat^2+bw^2); bw_res
                        <- alpha*res
            kfun <- function(h) {mean(dnorm(h,mean=bw_res,sd=bw))}; kfun
            <- Vectorize(kfun)
            out <- kfun(h1-h1bar)}
    # Finally, output the growth
    return(out)
}
###
## - Recruitment function
# Expected number of offspring
expected_offspring <- function(dens) {
    num_new <- nrow(seedlingdata[seedlingdata$dens==dens,])
    num_flow <- nrow(dat[dat$Stage2002==4 & dat $dens=_dens,])
    return(num_new/num_flow)
}
# Fit recruit size to a log-normal distribution
r_size <- fitdistr(seedlingdata$Height2003, "lognormal")$estimate
r_size_lo <- fitdistr (seedlingdata[seedlingdata$dens="Low",]$
    Height2003, "lognormal")$estimate
r_size_hi <- fitdistr(seedlingdata [seedlingdata$dens=" High" ,] $
    Height2003, "lognormal")$estimate
# Function for probability of recruit size
recruit_size <- function(zvec,dens) {
    if (dens="LLow") {my_dist <- r_size_lo}
    else if (dens="High") {my_dist <- r_size_hi}
    else {my_dist <- r_size}
    return(dlnorm(zvec, meanlog=my_dist["meanlog"], sdlog=my_dist["
        sdlog"], log=F))
}
```

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```
###
## - Kernel functions
# Survival-growth kernel
P_k <- function(h,meshpts, dens) {
    fxn <- function(z1,z, dens) {
        return(p_s(z,dens)*(1-p_r(z,dens)) *g_k(z1,z, dens))}
    return(h*(outer(meshpts,meshpts,fxn,dens)))}
# Reproduction kernel
R_k <- function(h,meshpts, dens,combine_repr) {
    fxn <- function(z1,z, dens) {
        return(p_s(z,dens)*p_r(z,dens)*expected_offspring(dens)*
                recruit_size(z1, combine_repr)/(h*sum(recruit_size (meshpts,
                combine_repr))))}
    return(h*(outer(meshpts,meshpts,fxn,dens)))}
###
## - Find kernels and population growth rates
# Function to do so
ipm_kernel <- function(h, meshpts, dens, combine_repr) {
    surv_growth <- P_k(h,meshpts, dens)
    fecundity <- R_k(h,meshpts, dens, combine_repr)
    kernel <- surv_growth + fecundity
    return(kernel)}
# Specify bounds, mesh, spacing
L<- 5
U <- 250
m<- 1000
h <- (U-L)/m
meshpts <- L + (1:m)*h - h/2
# Find kernels
```

277 k_lo <- ipm_kernel (h, meshpts, "Low","Low")
278

```
k_hi <- ipm_kernel(h,meshpts,"High","High")
# Find eigenvalues of kernels
lam_lo <- Re(eigen(k_lo,only.values=T)$values [1])
lam_hi <- Re(eigen(k_hi,only.values=T)$values[1])
# Print out the final values
lam_lo
lam_hi
###
## - Find eigenvectors
# Find eigenvectors of kernels
v_lo <- Re(eigen(t(k_lo))$vectors[,1])
v_hi<<-Re(eigen(t(k_hi))$vectors[,1])
w_lo <- Re(eigen(k_lo)$vectors[,1])
w_hi <- Re(eigen(k_hi)$vectors[,1])
# Normalize eigenvectors
v_lo <- v_lo/sum(v_lo)
v_hi<< v_hi/sum(v_hi)
w_lo <- w_lo/sum(w_lo)
w_hi <- w_hi/sum(w_hi)
```


## A. 2 Diagnostics

## A.2.1 Unintentional Eviction

```
###
## - Compute epsilons and rhos
# Functions to integrate over
not_evicted <- function(x,dens) {
```


\#\#\#
\#\# - Compute eviction at the stable size distribution
\# Probability of eviction below L
L_evicted $<-$ function (x, dens) \{
helper_fxn $<-$ function $(y)\left\{p_{-} s(x, d e n s) * g_{-}(y, x, d e n s)\right\}$
integrate (helper_fxn, $-\operatorname{Inf}, L)\}$
rho_lo_L<- function(x) \{L_evicted (x,"Low") \}
rho_hi_L $<-$ function (x) $\left\{L_{-}\right.$evicted (x,"High") \}
temp_lo_L <- lapply (meshpts, rho_lo_L)
temp_hi_L <- lapply (meshpts, rho_hi_L)
rho_lo_L_vec $<-\operatorname{array}($ as.numeric (unlist (lapply (1:m, function (i) \{ as
. numeric (temp_lo_L[[i]]\$value) \}))))
rho_hi_L_vec $<-\operatorname{array}($ as.numeric (unlist (lapply (1:m, function (i) \{as
. numeric (temp_hi_L[[i]]\$value) \}) )) )
\# Probability of eviction above U
U_evicted $<-$ function (x, dens) \{
helper_fxn $<-$ function $(y) \quad\left\{p_{-} s(x, d e n s) * g_{-}(y, x, d e n s)\right\}$
integrate (helper_fxn, U, Inf) $\}$
rho_lo_U <- function(x) \{U_evicted (x,"Low") \}
rho_hi_U $<-$ function (x) $\left\{U_{-}\right.$evicted (x, "High") \}
temp_lo_U <- lapply (meshpts, rho_lo_U)
temp_hi_U <- lapply (meshpts, rho_hi_U)
rho_lo_U_vec $<-$ array (as. numeric (unlist (lapply (1:m, function (i) \{as
. numeric (temp_lo_U[[i]]\$value) \}))))
rho_hi_U_vec $<-\operatorname{array}($ as. numeric (unlist (lapply (1:m, function (i) \{as
. numeric (temp_hi_U[[i]]\$value) \}))))
\# Unintentional eviction
format (w_lo \%*\% rho_lo_L_vec, scientific=F)
format (w_lo \%*\% rho_lo_U_vec, scientific=F)
w_hi $\% * \%$ rho_hi_L_vec
w_hi $\% * \%$ rho_hi_U_vec

| 76 |  |
| :---: | :---: |
| 77 | \#\#\# |
| 78 | \#\# -- Computing d lambda |
| 79 |  |
| 80 | \# Find d lambdas |
| 81 | inner_lo_L <- rho_lo_L_vec $\% * \% w_{\text {c }}$ lo |
| 82 | inner_hi_L <- rho_hi_L_vec $\% * \% w_{-}$hi |
| 83 |  |
| 84 | inner_lo_U <- rho_lo_U_vec $\% * \% w_{\text {_ }}$ lo |
| 85 | inner_hi_U <- rho_hi_U_vec $\% * \% w_{-}$hi |
| 86 |  |
| 87 | inner_lo <- v_lo \%** w_lo |
| 88 | inner_hi <- v_hi \%*\% w_hi |
| 89 |  |
| 90 | dlam_lo_L <- v_lo [1] * inner_lo_L / inner_lo |
| 91 | dlam_hi_L <- v_hi [1] * inner_hi_L / inner_hi |
| 92 | dlam_lo_U <- v_lo [m] * inner_lo_U / inner_lo |
| 93 | dlam_hi_U <- v_hi [m] * inner_hi_U / inner_hi |
| 94 |  |
| 95 | \# Print thems |
| 96 | dlam_lo_L |
| 97 | dlam_lo_U |
| 98 | dlam_hi_L |
| 99 | dlam_hi_U |
| 100 |  |
| 101 | dlam_lo_L + dlam_lo_U |
| 102 | dlam_hi_L + dlam_hi_U |

## A.2.2 Form of Growth

```
###
## - Table 2.2: Forms for growth (and first paragraph)
# Functions for the forms
quad_fun <- function(x, coefs) {
    a<-coefs [1]; b<-coefs[2]; return(a*x^2+b*x)}
```

quad3_fun $<-$ function (x, coefs) \{
$\mathrm{a}<-$ coefs [1]; $\mathrm{b}<-$ coefs [2]; $\mathrm{c}<-\operatorname{coefs}[3]$; return (a*x^2+b*x+c)\}
rick_fun $<-$ function (x, coefs) \{
$\mathrm{a}<-\operatorname{coefs}[1] ; \mathrm{b}<-\operatorname{coefs}[2] ;$ return $(a * x * \exp (-b * x))\}$
skel_fun $<-$ function (x, coefs) \{
$a<-\operatorname{coefs}[1] ; b<-\operatorname{coefs}[2] ;$ return $(a *(1-\exp (-b * x)))\}$
bvht_fun $<-$ function (x, coefs) \{
$a<-\operatorname{coefs}[1] ; \mathrm{b}<-\operatorname{coefs}[2] ;$ return $(a * x /(b+x))\}$
powr_fun $<-$ function (x, coefs) \{
$a<-\operatorname{coefs}[1] ; \mathrm{b}<-\operatorname{coefs}[2] ;$ return (a*x^b) \}
\# Fit the fits
dat_asymp <- dat_adult_low
quad_fit <- nls (h1 ~ quad_fun(h, c(a,b)), data=dat_asymp, start= list $(a=1, b=1)$ )
quad3_fit<- lm(h1~poly (h,2), data=dat_asymp)
rick_fit <- nls(h1 ~ rick_fun(h, c(a,b)), data=dat_asymp, start= list $(a=1, b=0.01))$
skel_fit <- nls(h1 ~ skel_fun(h, c(a,b)), data=dat_asymp, start= list $(a=100, b=0.01))$
bvht_fit <- nls(h1 ~ bvht_fun(h, c(a,b)), data=dat_asymp, start= list $(a=1, b=1)$ )
powr_fit <- nls(h1 ~ powr_fun(h, c(a,b)), data=dat_asymp, start= list $(a=1, b=1)$ )
\# AIC for the above fits
AIC (grow_fit_low)
AIC (quad _fit)
AIC(quad3_fit)
AIC(rick_fit)
AIC(skel_fit)
AIC(bvht_fit)
AIC (powr_fit)
\# Fit some more fits
dat_asymp <- dat_adult_high
quad_fit <- nls (h1 ~ quad_fun(h, c(a,b)), data=dat_asymp, start= list $(a=1, b=1)$ )

```
quad3_fit<- lm(h1~poly (h, 2), data=dat_asymp)
rick_fit <- nls(h1 ~ rick_fun(h, c(a,b)), data=dat_asymp, start=
    list \((a=1, b=0.01)\) )
skel_fit <- nls(h1~ skel_fun(h, c(a,b)), data=dat_asymp, start=
    list \((a=100, b=0.01))\)
bvht_fit <- nls(h1 ~ bvht_fun(h, c(a,b)), data=dat_asymp, start=
    list \((a=1, b=1)\) )
powr_fit <- nls(h1 ~ powr_fun(h, c(a,b)), data=dat_asymp, start=
    list \((a=1, b=1)\) )
\# AIC for even more fits
AIC (grow_fit_high)
AIC (quad_fit)
AIC(quad3_fit)
AIC(rick_fit)
AIC(skel_fit)
AIC(bvht_fit)
AIC (powr_fit)
```


## A.2.3 Normality of Residuals and Homoskedasticity

```
###
## - Normality of residuals and heteroskedasticity (second
    growth paragraph)
# Low dens diagnostics
plot(dat_adult_low$h, dat_adult_low$h1)
plot (dat_adult_low $h, grow_fit_low$residuals)
plot(dat_adult_low $h, hat(model.matrix (grow_fit_low)))
plot(dat_adult_low$h, cooks.distance(grow_fit_low ))
qqnorm(grow_fit_low$residuals)
shapiro.test(grow_fit_low$residuals)
bptest(grow_fit_low)
AIC(grow_fit_low)
# High dens diagnostics
```

```
plot(dat_adult_high$h,dat_adult_high$h1)
plot(dat_adult_high$h,grow_fit_high$residuals)
plot(dat_adult_high$h,hat(model.matrix (grow_fit_high)))
plot(dat_adult_high$h, cooks.distance(grow_fit_high))
qqnorm(grow_fit_high$residuals)
shapiro.test(grow_fit_high$residuals)
bptest(grow_fit_high)
AIC(grow_fit_high)
```


## A.2.4 Log-normality of the Recruit Size Distribution

```
###
## - Test fit of recruit dist'n
df <- seedlingdata$h1
df_lo <- seedlingdata[seedlingdata$dens="Low",] $h1
df_hi<- seedlingdata[seedlingdata$dens=" High",] $h1
t.test(log(df_lo), log(df_hi), paired=F,var.equal=F, conf.level
    =0.95)
shapiro.test(log(df))
shapiro.test(log(df_lo))
shapiro.test(log(df_hi))
sd(log(df_lo))
sd(log(df_hi))
```


## A. 3 Bootstrapping

## A.3.1 Bootstrap population growth rate

1 2 3 4
bootstrap_helper $<-$ function (adult_dat, seedling_dat) \{
\#\#\# - The data set to perform regression on

```
# Which subset of data to use
dat <- adult_dat
# Use this one for growth
#dat_adult <- dat[dat$Stage2003!=1 & dat$Stage2003!=4 & dat$
    Stage2003!=5,]
dat_adult <- dat[dat$Stage2003!=4 & dat$Stage2003!=5,]
####
## -- Survival and reproduction
# Logistic regression
surv_fit <- glm(survived ~h, data=dat, family="binomial")
repr_fit <- glm(reproduced ~h, data=dat, family="binomial")
# Response functions
p_s <- function(h){return(predict(surv_fit, data.frame(h=h),
    type="response"))}
p_r <- function(h){return(predict(repr_fit, data.frame(h=h),
    type="response"))}
```

```
###
```


### 

## -- Growth

## -- Growth

# Fit growth

# Fit growth

grow_fit <- lm(h1~h, data=dat_adult)
grow_fit <- lm(h1~h, data=dat_adult)

# Growth kernel

# Growth kernel

g_k <- function(h1,h) {
g_k <- function(h1,h) {
h1bar <- predict(grow_fit, newdata=data.frame(h=h), type="
h1bar <- predict(grow_fit, newdata=data.frame(h=h), type="
response")
response")
sdhat <- sqrt(sum(residuals(grow_fit)^2)/df.residual(grow_fit
sdhat <- sqrt(sum(residuals(grow_fit)^2)/df.residual(grow_fit
))
))
return(dnorm(h1, mean=h1bar, sd=sdhat))}
return(dnorm(h1, mean=h1bar, sd=sdhat))}

### 

### 

## -- Recruitment kernel

```
## -- Recruitment kernel
```

\# Expected number of offspring (num_new / num_flowering)
expected_offspring <- nrow (seedling_dat)/nrow (dat[dat\$Stage2002
$==4$,$] )$
\# Fit recruit size to a log-normal distribution
$r_{\text {_ }}$ size $<-$ fitdistr (seedling_dat $\$$ Height2003, "lognormal") $\$$
estimate
\# Function for probability of recruit size
recruit_size $<-$ function (zvec) \{return (dlnorm(zvec, meanlog=r_
size["meanlog"], sdlog=r_size["sdlog"], log=F)) \}
\#\#\#
\#\# - Implement kernels
\# Survival-growth
P_k <- function(h, meshpts) \{

return (h* (outer (meshpts, meshpts, fxn)) ) \}
\# Reproduction
R_k $<-$ function (h, meshpts, dens, combine_repr) \{
fxn $<-$ function $(z 1, z)$ \{return (p_s $(z) * p_{-} r(z) * \operatorname{expected}$
offspring*recruit_size(z1)) \}
return (h* (outer (meshpts, meshpts, fxn) ) ) \}
\# Full kernel
ipm_kernel $<-$ function(h, meshpts) \{return (P_k(h, meshpts) $+\mathrm{R}_{-} \mathrm{k}(\mathrm{h}$,
meshpts) ) \}

```
###
## -- Find growth rate
# Specify bounds, mesh, spacing
L<- 5
U<- 250
m}<-10
```

| 71 | $\mathrm{h}<-(\mathrm{U}-\mathrm{L}) / \mathrm{m}$ |
| :---: | :---: |
| 72 | meshpts $<-\mathrm{L}+(1: \mathrm{m}) * \mathrm{~h}-\mathrm{h} / 2$ |
| 73 |  |
| 74 | \# Determine kernels and eigenvectors |
| 75 | $\mathrm{P}<-\mathrm{P}$ - $\mathrm{k}(\mathrm{h}, \mathrm{mesh} \mathrm{ts})$ |
| 76 | $\mathrm{R}<-\mathrm{R}_{-} \mathrm{k}$ (h, meshpts) |
| 77 | kernel $<-\mathrm{P}+\mathrm{R}$ |
| 78 | fund_mat $<-$ R \%*\% solve (diag (m)-P) |
| 79 |  |
| 80 | $\mathrm{w}<-\operatorname{Re}($ eigen (kernel)\$vectors $[, 1])$ |
| 81 | $\mathrm{w}<-\mathrm{w} / \operatorname{sum}(\mathrm{w})$ |
| 82 |  |
| 83 | \# Determine values to report |
| 84 | height <- w \%*\% meshpts |
| 85 | lambda <- Re(eigen (kernel) \$values [1]) |
| 86 | R0 <- Re(eigen (fund_mat) \$values [1]) |
| 87 | Tval $<-\log (\mathrm{R} 0) / \log (\mathrm{lambda})$ |
| 88 | return (list ("height"=height, "lambda"=lambda, "R0"=R0," Tval"=Tval )) |
| 89 | $\}$ 边 |
| 90 |  |
| 91 |  |
| 92 | \#\#\# |
| 93 | \#\# - Function to verify if subset is ok for bootstrapping |
| 94 |  |
| 95 | verify _sample <- function (dat) \{ |
| 96 | \# By default it should be fine, I just need to check if there might be a problem |
| 97 | out $<-0$ |
| 98 | \# These subsets will help, I remove "NA's" |
| 99 | seed_dat <- dat[dat\$class="Seedling", ] |
| 100 | adlt_dat <- dat[dat \$class! = "Seedling" \& ! is.na(dat \$ survived) , ] |
| 101 | \# More helpful subsets: first digit is survival, second is reproduction |
| 102 | sr_dat_ $00<-$ adlt_dat[adlt_dat $\$$ survived $==0$, ] |
| 103 | sr_dat_10<- adlt_dat[adlt_dat\$survived==1 \& adlt_dat $\$$ reproduced $==0$, |

104

```
    sr_dat_11<- adlt_dat[adlt_dat$survived==1 & adlt_dat$
            reproduced==1,]
    # Need to have enough seedlings to parameterize
    if (nrow(seed_dat)<=2) {out <- NA}
    # Return whether the data set is ok or not
    return(out)
}
###
## - Function for bootstrapping
bootstrap_main <- function(dat, sample_size, N) {
    # Number of rows to sample from
    nr <- nrow(dat)
    # Vector to store output
    heights <- array (NA, dim=c(N))
    lambdas <- array (NA, dim=c (N))
    R0s <- array (NA, dim=c (N))
    Ts <- array (NA, dim=c(N))
    # Index
    i = 1
    j = 0
    # Bootstrap step
    while (i <= N) {
        # Randomly sample from data set
        indices <- sample.int(nr, sample_size, replace=T)
        dat_new <- dat[indices,]
        verf <- verify _sample(dat_new)
        # Check to see if data will work
        if (!is.na(verf)) {
            out <- bootstrap_helper(dat_new[dat_new$class!="Seedling"
            ,], dat_new[dat_new$class="Seedling",])
            heights[i] <- out$height
            lambdas[i] <- out$lambda
            R0s[i] <- out$R0
            Ts[i] <- out$Tval
            print(i)
```

```
        i <- i+1
        }
        else {
        print(c(i,"bad"))
        j <- j+1
        }
    }
        print(c("There\iotawere",j," errors"))
        return(list("heights"=heights,"lambdas"=lambdas,"R0s"=R0s,"Ts"=
            Ts))
}
###
## - Bootstrap
# Subset data into high and low
dat_lo <- hmdata_no_outlier[hmdata_no_outlier $dens="Low",]
dat_hi <- hmdata_no_outlier [hmdata_no_outlier$dens=" High",]
# Compute the bootstrap confidence intervals
boot_distn_lo <- bootstrap_main(dat_lo, nrow(dat_lo),5000)
boot_distn_hi <- bootstrap_main(dat_lo , nrow(dat_hi),5000)
# Break into constituent parts
h_distn_lo <- boot_distn_lo$heights
h_distn_hi <- boot_distn_hi$heights
lam_distn_lo <- boot_distn_lo$lambdas
lam_distn_hi <- boot_distn_hi$lambdas
# Ensure each value is valid
length(lam_distn_lo[lam_distn_lo<0])
length(lam_distn_hi[lam_distn_hi<0])
length(lam_distn_lo[is.na(lam_distn_lo)])
length(lam_distn_hi[is.na(lam_distn_hi)])
# Bootstrap means
```

```
h_boot_mu_lo <- mean(h_distn_lo)
h_boot_mu_hi <- mean(h_distn_hi)
lam_boot_mu_lo <- mean(lam_distn_lo)
lam_boot_mu_hi <- mean(lam_distn_hi)
# Bootstrap confidence intervals
h_boot_ci_hi<- quantile(h_distn_hi, probs=c (0.025,0.975))
h_boot_ci_lo <- quantile(h_distn_lo, probs=c (0.025,0.975))
lam_boot_ci_hi<- quantile(lam_distn_hi, probs=c (0.025,0.975))
lam_boot_ci_lo<- quantile(lam_distn_lo, probs=c (0.025,0.975))
###
## - Height table
# Using normal bootstrapping sample size
w_lo %*% meshpts
h_boot_mu_lo
h_boot_ci_lo
w_hi %*% meshpts
h_boot_mu_hi
h_boot_ci_hi
###
## -- Population growth rate table
# Using normal bootstrapping sample size
lam_lo
lam_boot_mu_lo
lam_boot_ci_lo
lam_hi
lam_boot_mu_hi
lam_boot_ci_hi
```

| 215 |  |
| :---: | :---: |
| 217 | \#\# |
| 218 | \#\# -- Upsampling data |
| 219 |  |
| 220 | \# Bootstrap |
| 221 | boot_distn_lo_100<- bootstrap_main (dat_lo, 100, 5000) |
| 222 | boot_distn_lo_200<- bootstrap_main (dat_lo, 200, 5000) |
| 223 | boot_distn_lo_300<- bootstrap_main (dat_lo, 300,5000) |
| 224 | boot_distn_lo_ $400<-$ bootstrap_main (dat_lo, 400,5000) |
| 225 | boot_distn_lo_500<- bootstrap_main (dat_lo, 500,5000) |
| 226 | boot_distn_lo_600<- bootstrap_main (dat_lo, 600,5000) |
| 227 |  |
| 228 | boot_distn_hi_100<- bootstrap_main (dat_hi , 100, 5000) |
| 229 | boot_distn_hi_ $200<-$ bootstrap_main (dat_hi , 200, 5000) |
| 230 | boot_distn_hi_ $300<-$ bootstrap_main (dat_hi , 300, 5000) |
| 231 | boot_distn_hi_ $400<-$ bootstrap_main (dat_hi , 400, 5000) |
| 232 | boot_distn_hi_ $500<-$ bootstrap_main (dat_hi , 500, 5000) |
| 233 | boot_distn_hi_600<- bootstrap_main (dat_hi, 600,5000) |
| 234 |  |
| 235 | \# Give name to bootstrap distributions |
| 236 | lam_distn_lo_ $100<-$ boot_distn_lo_ 100 \$lambdas |
| 237 | lam_distn_lo_ $200<-$ boot_distn_lo_ 200 \$ lambdas |
| 238 | lam_distn_lo_ $300<-$ boot_distn_lo_ 300 \$ lambdas |
| 239 | lam_distn_lo_ $400<-$ boot_distn_lo_ 400 \$ lambdas |
| 240 | lam_distn_lo_500<- boot_distn_lo_ 500 \$lambdas |
| 241 | lam_distn_lo_600<- boot_distn_lo_ 600 \$ lambdas |
| 242 |  |
| 243 | lam_distn_hi_ $100<-$ boot_distn_hi_ 100 \$lambdas |
| 244 | lam_distn_hi_ $200<-$ boot_distn_hi_ 200 \$lambdas |
| 245 | lam_distn_hi_ $300<-$ boot_distn_hi_ 300 \$lambdas |
| 246 |  |
| 247 | lam_distn_hi_ $500<-$ boot_distn_hi_ 500 \$lambdas |
| 248 | lam_distn_hi_600<- boot_distn_hi_600\$lambdas |
| 249 |  |
| 250 | \# Compute bootstrap means |
| 251 | lam_ boot_mu_lo_ $100<-$ mean (lam_distn_lo_ 100) |
| 252 | lam_boot_mu_lo_200<- mean (lam_distn_lo_ 200) |

253

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lam_boot_mu_lo_300<- mean(lam_distn_lo_300)
lam_boot_mu_lo_ $400<-$ mean (lam_distn_lo_400)
lam_boot_mu_lo_500<- mean(lam_distn_lo_500)
lam_boot_mu_lo_600<- mean (lam_distn_lo_600)
lam_boot_mu_hi_100<- mean(lam_distn_hi_100)
lam_boot_mu_hi_200<- mean(lam_distn_hi_200)
lam_boot_mu_hi_300<- mean (lam_distn_hi_300)
lam_boot_mu_hi_400<- mean(lam_distn_hi_400)
lam_boot_mu_hi_500<- mean(lam_distn_hi_500)
lam_boot_mu_hi_ $600<-$ mean (lam_distn_hi_600)
\# Compute bootstrap CIs
lam_boot_ci_lo_100<- quantile (lam_distn_lo_100, probs=c (0.025, 0.975) )
lam_boot_ci_lo_ $200<-$ quantile (lam_distn_lo_ 200 , probs=c (0.025, 0.975) )
lam_boot_ci_lo_300<- quantile (lam_distn_lo_300, probs=c (0.025, 0.975) )
lam_boot_ci_lo_ $400<-$ quantile (lam_distn_lo_ 400 , probs=c (0.025, 0.975) )
lam_boot_ci_lo_500<- quantile (lam_distn_lo_500, probs=c (0.025, 0.975) )
lam_boot_ci_lo_600<- quantile (lam_distn_lo_600, probs=c $(0.025,0.975))$
lam_boot_ci_hi_100<- quantile (lam_distn_hi_100, probs=c (0.025, 0.975) )
lam_boot_ci_hi_200<- quantile(lam_distn_hi_200, probs=c (0.025, 0.975) )
lam_boot_ci_hi_300<- quantile (lam_distn_hi_300, probs=c (0.025, 0.975) )
lam_boot_ci_hi_ $400<-$ quantile (lam_distn_hi_ 400 , probs=c ( $0.025,0.975$ ) )
lam_boot_ci_hi_500<- quantile(lam_distn_hi_500, probs=c (0.025, 0.975))
lam_boot_ci_hi_600<- quantile (lam_distn_hi_600, probs=c (0.025, 0.975) )

| 279 |  |
| :---: | :---: |
| 280 | \# Print values |
| 281 | lam_boot _mu_lo _ 100 |
| 282 | lam_boot_ci_lo_100 |
| 83 |  |
| 284 | lam_ boot_mu_ lo _ 200 |
| 285 | lam_boot_ci_lo_200 |
| 286 |  |
| 287 | lam_ boot _mu_lo _ 300 |
| 288 | lam_boot_ci_lo_300 |
| 289 |  |
| 290 | lam_ boot _mu_ lo _ 400 |
| 291 | lam_ boot_ ci _ lo _ 400 |
| 292 |  |
| 293 | lam_boot _mu_lo _ 500 |
| 294 | lam_boot_ci_lo_500 |
| 295 |  |
| 296 | lam_ boot_mu_ lo _ 600 |
| 297 | lam_ boot_ci_lo_600 |
| 298 |  |
| 299 |  |
| 300 | lam_ boot_mu_hi_ 100 |
| 301 | lam_boot_ci_hi_100 |
| 302 |  |
| 303 | lam_ boot_mu_hi_ 200 |
| 304 | lam_boot_ci_hi_200 |
| 305 |  |
| 306 | lam_boot _mu_hi_300 |
| 307 | lam_boot_ci_hi_300 |
| 308 |  |
| 309 | lam_ boot_mu_hi _ 400 |
| 310 | lam_boot_ci_hi_400 |
| 311 |  |
| 312 | lam_boot_mu_hi_500 |
| 313 | lam_boot_ci_hi_500 |
| 314 |  |
| 315 | lam_ boot_mu_hi_600 |
| 316 | lam_boot_ci_hi_600 |

## A. 4 Methods Plots

## A.4.1 Plot vital rate functions

```
###
## - Plot vital rates
# Specify data set to use
dat <- adultdata_no_outlier
# Real data to plot (ticks or points as necessary)
# Survival ticks
dat_surv <- dat[!is.na(dat$survived),]
dat_surv_lo <- dat_surv[dat_surv$dens="Low",]
dat_surv_hi <- dat_surv[dat_surv$dens=" High",]
# Reproduction ticks
dat_repr <- dat[!is.na(dat$reproduced) ,]
dat_repr_lo <- dat_repr[dat_repr$dens="Low" ,]
dat_repr_hi <- dat_repr[dat_repr$dens=" High" ,]
# Growth points
dat_veg <- dat[dat$Stage2003!=4 & dat$Stage2003!=5 & !is.na(dat$h
    ),]
dat_veg_lo <- dat_veg[dat_veg$dens="Low",]
dat_veg_hi <- dat_veg[dat_veg$dens=" High",]
# Make lines that will be plotted
xmax_lo <- 150
xmax_hi <- 220
x_lo <- seq (0,xmax_lo, length=100)
x_hi}<-\operatorname{seq}(0,xmax_hi, length=100
s_lo<- p_s(x_lo,"Low")
s_hi<< p_s(x_hi,"High")
```

30
31
32
33
$r_{-}$lo $<-p_{-} r\left(x_{-}\right.$lo, "Low" $)$
$r_{-} h i<-p_{-} r\left(x_{-} h i, " H i g h "\right)$
\# PLOTS
par (mfrow=c $(3,2))$
dx $<-0.03$
dy $<-1$
line_frac <-1
scale_text <-1.3


cex. lab=scale_text, cex. axis=scale_text)
points (dat_surv_lo\$h, dat_surv_lo\$survived, pch="|", cex=line_frac)
mtext (" (a)", side=3, adj=dx, padj=dy, line=-1.3)
plot (x_hi, s_hi, type="l", lwd=2, xlim=c (0,xmax_hi), ylim=c (0,1),

cex.lab=scale_text, cex. axis=scale_text)
points (dat_surv_hi\$h, dat_surv_hi\$survived, pch="|", cex=line_frac)
mtext (" (b)", side=3, adj=dx, padj=dy, line=-1.3)
plot (x_lo, r_lo, type=" $1 ", \operatorname{lwd}=2, x \lim =c\left(0, x \max \_l o\right), y l i m=c(0,1)$,
xlab="Height of $\lrcorner$ plant $\lrcorner(\mathrm{cm})$ ", ylab="Probability $\lrcorner$ of $\lrcorner$ flowering",
cex. lab=scale_text, cex. axis=scale_text)
points (dat_repr_lo\$h, dat_repr_lo\$reproduced, pch="|", cex=line_frac
)
mtext (" (c)", side=3, adj=dx, padj=dy, line=-1.3)
plot (x_hi, r_hi, type="l", lwd=2, xlim=c (0,xmax_hi), ylim=c (0,1),

cex. lab=scale_text, cex.axis=scale_text)
points (dat_repr_hi\$h, dat_repr_hi\$reproduced, pch="|", cex=line_frac
)
mtext (" (d)", side=3, adj=dx, padj=dy, line=-1.3)
plot (dat_veg_lo\$h, dat_veg_lo\$h1, pch=16, xlim=c(0,xmax_lo),ylim=c
$(0,220)$,

```
60 xlab="Height」in」2002」(cm)", ylab=" Heightьin \(2003 \sqcup(\mathrm{~cm})\) ", cex.
    lab=scale_text, cex. axis=scale_text)
    segments (min (dat_veg_lo\$h), y0=as.numeric (predict (grow_fit_low,
        data.frame (h=min(dat_veg_lo\$h))),
        \(\max (\) dat_veg_lo \(\$ \mathrm{~h}), \mathrm{y} 1=\) as.numeric (predict (grow_fit_low,
                data.frame (h=max (dat_veg_lo\$h))), lwd=2)
mtext (" (e)", side=3, adj=dx, padj=dy, line=-1.3)
plot (dat_veg_hi\$h, dat_veg_hi\$h1, pch=16, xlim=c (0,xmax_hi),ylim=c
        (0,220) ,
            xlab=" Heightıin 2002\(\lrcorner(\mathrm{cm})\) ", ylab=" Heightıin 2003\(\lrcorner(\mathrm{cm})\) ", cex.
            lab=scale_text, cex. axis=scale_text)
segments (min (dat_veg_hi\$h), y0=as.numeric (predict (grow_fit_high,
    data.frame (h=min(dat_veg_hi\$h)))),
```



```
                data.frame (h=max (dat_veg_hi\$h)))), lwd=2)
mtext(" (f)", side=3,adj=dx, padj=dy, line=-1.3)
\(\operatorname{par}(\operatorname{mfrow}=c(1,1))\)
```


## A．4．2 Plot recruit size distribution

```
###
## - Plot recruit size distn
# Specify data set to use
dat <- seedlingdata_no_outlier
dat_recr_lo <- dat[dat$dens="Low" ,] $Height2003
dat_recr_hi <- dat[dat$dens=" High" ,] $Height2003
# Make lines to plot
xmin_lo <- min(dat_recr_lo)
xmax_lo <- max(dat_recr_lo)
x_lo <- seq(xmin_lo, xmax_lo, length=100)
y_lo <- recruit_size(x_lo,"Low")
xmin_hi <- min(dat_recr_hi)
```

```
xmax_hi <- max(dat_recr_hi)
x_hi <- seq(xmin_hi, xmax_hi, length=100)
y_hi <- recruit_size(x_hi,"High")
# PLOTS
par(mfrow=c (1,2))
distnce <- 0.02 # -0.22
scale_text <- 1
hist(dat_recr_lo, breaks=7, xlim=c(xmin_lo - 5,xmax_lo) , ylim=c
    (0,0.05),
```



```
        cex.lab=scale_text, cex.axis=scale_text, main="", freq=FALSE)
lines(x_lo,y_lo ,type=" l",lwd=2)
mtext("(a)", side=3,adj=distnce, line=-1.3)
hist(dat_recr_hi, breaks=9, xlim=c(xmin_hi-2,xmax_hi),ylim=c
    (0,0.02),
        xlab="Recruit^heightヶ(cm)", ylab="Proportion\_of recruits",
        cex.lab=scale_text, cex.axis=scale_text, main="", freq=FALSE)
lines(x_hi,y_hi,type=" l",lwd=2)
mtext("(b)", side=3,adj=distnce, line=-1.3)
par(mfrow=c(1,1))
```


## A. 5 Results Plots

## A.5.1 Plot predicted size distribution vs observations

```
# Main data frames
dat_a<- adultdata_no_outlier[adultdata_no_outlier $Stage2003!=4,]
dat_s <- seedlingdata_no_outlier
# Neat little data frames
adlt_lo <- dat_a[dat_a$dens=" Low",]
adlt_hi <- dat_a[dat_a$dens=" High",]
seed_lo <- dat_s[dat_s$dens=" Low",]
seed_hi <- dat_s[dat_s$dens=" High",]
```

```
# Isolate heights w/o NAs
h_adlt_lo <- adlt_lo$h1[!is.na(adlt_lo$h1)]
h_seed_lo <- seed_lo$h1[!is.na(seed_lo$h1)]
h_adlt_hi <- adlt_hi$h1[!is.na(adlt_hi$h1)]
h_seed_hi <- seed_hi$h1[!is.na(seed_hi$h1)]
# Combine adult and seedling data
observed_lo <- c(h_adlt_lo,h_seed_lo)
observed_hi <- c(h_adlt_hi,h_seed_hi)
# Predicted values
predicted_lo <- w_lo/h
predicted_hi <- w_hi/h
# Mean heights
mean_h_lo_prd <- w_lo %*% meshpts
mean_h_hi_prd <- w_hi %*% meshpts
mean_h_lo_obs <- mean(observed_lo)
mean_h_hi_obs <- mean(observed_hi)
mean_h_lo_obs
mean_h_lo_prd
mean_h_hi _obs
mean_h_hi_prd
# Values are found by running 'ch2_bootstrapping.R' and added
    manually
# I do not do this programmatically since running the
# bootstrap file every time I want to make this fig
boot_summary_lo <- c(86.65732, 94.60086, 102.26890)
boot_summary_hi <- c(67.99328, 73.05466, 78.14434)
# PLOTS
par (mfrow=c(1,2))
distnce <- 0.02
scale_text <- 1
hist (observed_lo, breaks=10,freq=F, xlim=c(0,U),
    xlab="Height", ylab=" Proportion",main="",
```

```
        cex.lab=scale_text, cex.axis=scale_text)
lines(meshpts, predicted_lo , lwd=2)
mtext("(a)", side=3,adj=distnce, line=-1.3)
hist(observed_hi, breaks=10, freq=F, xlim=c(0,U),
    xlab="Height", ylab="Proportion",main="",
    cex.lab=scale_text, cex.axis=scale_text)
lines(meshpts, predicted_hi, lwd=2)
mtext("(b)", side=3,adj=distnce, line= - 1.3)
par(mfrow=c(1,1))
```


## Appendix B

## Chapter 3 Code

## B. 1 Simulation Function

|  | \# Import libraries |
| :---: | :---: |
| 2 <br> 3 <br> 4 <br> 5 <br> 6 <br> 7 <br> 8 <br> 9 <br> 10 <br> 11 <br> 12 <br> 13 <br> 14 <br> 15 <br> 16 <br> 17 <br> 18 <br> 19 <br> 20 | library (lme4) \# for mixed-effects modelling |
|  | library (MASS) \# for fitting negative binomials |
|  | library(lmtest) \# for performing likelihood ratio tests |
|  |  |
|  |  |
|  | \#\#\# |
|  | \#\# - Read in the data |
|  |  |
|  | \# The same data as was used in Chapter 2 |
|  | hmdata <- read.csv("./data/HuelsNewData.csv", na.strings=" "") |
|  |  |
|  | \# The number of seedlings per plot |
|  | $\begin{aligned} & \text { seedling numbers }<- \text { read.csv("./data/HuelsSeedlingData.csv", na. } \\ & \text { strings=" ") } \end{aligned}$ |
|  |  |
|  |  |
|  | \#\#\# - |
|  | \#\# - Define state variable |
|  |  |
| 20 | \# I use height as the state variable |

```
hmdata["h"] <- hmdata [["Height2002"]]
hmdata ["h1"] <- hmdata [["Height2003"]]
\#\#\#
\#\# - Create columns for site and plot
\# The site is the first letter of the ID
hmdata["site"] <- substr(hmdata\$ID,1,1) \# first letter of
    individual ID
\# Determine the plot number
hmdata["plot"] <- substr(hmdata\$ID,1, as.numeric (lapply (as.
    character (hmdata\$ID), nchar)) - 2 )
\#\#\#
\#\# - Remove bad data
\# Remove Allendorf data due to grazing damage
hmdata <- hmdata[hmdata\$site! ="A" ,]
\#\#\#
\#\# - Find biomass in each plot
\# NOTES:
\# mass_exp: "Global allocation rules for patterns of biomass
    partitioning in seed plants"
\# I divide by 100 to convert to meters (so the numbers are a
    reasonable size)
\# Estimate biomass of each plant
euclid <- \(3 \quad \#\) exponent=1
wbe \(<-4 \quad \#\) one-fourth law of West, Brown, and Enquist
otl \(<-9 / 2 \quad \#\) one-third law
upper_exp \(<-6 \#\) max value
mass_exp <- wbe
```

55
hmdata ["mass"] <- ifelse(!is.na(hmdata\$h), (hmdata["h"]/100) ^mass
_ exp,0)
hmdata["mass1"] $<-$ ifelse (!is.na(hmdata\$h1),(hmdata["h1"]/100)^
mass_exp, 0 )
\# Find the density of each plot based on biomass
dens_df <- aggregate (mass ~ plot, hmdata, sum)
names (dens_df)[names (dens_df)="mass"] <-"dens"
dens1_df <- aggregate(mass1 ~ plot, hmdata, sum)
names (dens1_df)[names (dens1_df)="mass1"] <-"dens1"
\# Merging the data frames to get seedlings/plot and biomass/plot
hmdata $<-$ merge(hmdata, seedling_numbers, by="plot")
hmdata <- merge(hmdata, dens_df, by="plot")
hmdata <- merge(hmdata, dens1_df, by="plot")
\#\#\#
\#\# - Survival, reproduction, and classification (i.e. seedling
or adult)
\# 1 means survived, 0 means did not survive, NA means flowered (
so of course it's dead)
hmdata["survived"] <- ifelse (hmdata[["Stage2002"]] $=4$, NA,
ifelse(hmdata [["Stage2003"]] =5, 0, 1))
\# 1 means reproduced, 0 means did not reproduce, NA means
individual died without reproducing
hmdata ["reproduced"] <- ifelse (hmdata [["Stage2003"]] =4, 1,
ifelse (hmdata [["Stage2003"]] $=5$, NA, 0) )
\# If a plant was recorded in 2003 but not 2002 it was a seedling
in 2002
hmdata["class"] <- ifelse(is.na(hmdata[["Stage2002"]]), "Seedling
", "Adult")

| 84 | \#\#\# |
| :---: | :---: |
| 85 | \#\# - Remove outliers |
| 86 |  |
| 87 | \# Remove outliers (has to be done AFTER mass calcs) |
| 88 | hmdata <- hmdata[hmdata\$ID ! =" KXIII09" \& hmdata\$ID ! ="VCI04" , ] |
| 89 |  |
| 90 | \# Remove a reproductive outlier |
| 91 | hmdata $<-$ hmdata[hmdata $\$$ reproduced $==0 \mid$ hmdata $\$ h 1>100 \mid$ is.na $($ hmdata\$reproduced) ,] |
| 92 |  |
| 93 |  |
| 94 | \#\#\# |
| 95 | \#\# - - Split single data frame into one for seedlings and one for adult plants |
| 96 |  |
| 97 | seedlingdata <- hmdata [hmdata ${ }^{\text {c class }}$ - "Seedling", ] |
| 98 | hmdata $<-$ hmdata[hmdata\$class!="Seedling",] |
| 99 |  |
| 100 |  |
| 101 | \#\#\# |
| 102 | \#\# - The data set to perform regression on |
| 103 |  |
| 104 | \# Which subset of data to use |
| 105 | dat $<-$ hmdata |
| 106 |  |
| 107 |  |
| 108 | \#\#\# |
| 109 | \#\# - Probability of survival |
| 110 |  |
| 111 | \# Logistic regression for survival probability depending on leaf stem diameter |
| 112 | surv_fit <- glm(survived ${ }^{\text {~ }} \mathrm{h}$, data=dat, family="binomial") |
| 113 |  |
| 114 | \# Survival function |
| 115 | p_s <- function (h) \{ |
| 116 | tempdf $<-$ data.frame (h=h) |
| 117 | return(predict(surv_fit, tempdf, type="response")) |
| 118 |  |

```
1 1 9
```

120
\#\#\#
\#\# - Probability of reproduction
\# Logistic regression for probability of reproduction depending on leaf stem diameter
repr_fit_rl $<-$ glm (reproduced $\sim^{\sim} h$, data=dat, family="binomial") repr_fit_ac $<-$ glm(reproduced $\sim_{~ h+d e n s, ~ d a t a=d a t, ~ f a m i l y=" ~}^{\text {" }}$ binomial")
\# Reproduction function
\# I have two cases since I have two models
p_r $<-$ function (h, dens, modl) \{
dens $<-$ as.numeric (dens)
if (modl="rl") \{ tempdf $<-$ data.frame $(\mathrm{h}=\mathrm{h})$ my_fit <-repr_fit_rl\}
else if (modl=" ac") \{ tempdf <- data.frame (h=h, dens=dens)
my_fit <- repr_fit_ac $\}$
return(predict(my_fit, tempdf, type="response")) \}
\#\#\#
\#\# - Expected size increase (non-flowering plants)
\# I need to exclude flowering plants and dead plants...
hmdata_adult $<-$ hmdata[hmdata $\$$ Stage $2003!=4$ \& hmdata $\$$ Stage2003!= 5 ,]
\# ... and don't forget seedlings
hmdata_adult $<-$ hmdata_adult [!is.na(hmdata_adult \$ Height2002), ]
\# Fit growth using a linear model
grow_fit_a $<-\operatorname{lm}\left(h 1^{\sim} h\right.$, data=hmdata_adult)
\# Define growth kernel

```
g_a<< function(h1,h) {
    newdata <- data.frame(h=h)
    h1bar <- predict(grow_fit_a, newdata=newdata, type="response"
        )
    res <- residuals(grow_fit_a)
    df.res <- df.residual(grow_fit_a)
    sse <- sum(res ^2)
    sdhat <- sqrt(sse/df.res)
    return(dnorm(h1, mean=h1bar, sd=sdhat))
}
###
## - Expected size increase (flowering plants)
# I only include flowering plants
hmdata_flow <- hmdata[hmdata $ Stage2003==4,]
# Fit growth using a linear model
grow_fit_r <- lm(h1~h, data=hmdata_flow )
# Define growth kernel
g_r<< function(h1,h) {
    newdata <- data.frame(h=h)
    h1bar <- predict(grow_fit_r, newdata=newdata, type="response"
        )
    res <- residuals(grow_fit_r)
    df.res <- df.residual(grow_fit_r)
    sse <- sum(res ` 2)
    sdhat <- sqrt(sse/df.res)
    return(dnorm(h1, mean=h1bar, sd=sdhat))
}
###
## - Number of recruits
```

```
# Function for probability of recruit size
recruit_size <- function(zvec) {
    return(dlnorm(zvec, meanlog=r_size["meanlog"], sdlog=r_size["
        sdlog"], log=F))
}
###
## - Implement kernels
# Survival-growth
P_k <- function(h,meshpts, dens,modl) {
    fxn <- function(z1,z, params) {
        dens <- params[1]; modl <- params[2]
        return(p_s (z)*(1-p_r(z,dens,modl))*g_a(z1, z ) )}
    params <- c(dens,modl)
    return(h*(outer(meshpts,meshpts,fxn, params)))}
# Reproduction
R_k <- function(h,meshpts, dens,modl) {
    fxn <- function(z1,z, params) {
        dens <- params[1]; modl <- params[2]
        return(p_s(z)*p_r(z, dens,modl)*g_r(z1,z))}
    params <- c(dens,modl)
    return(h*(outer(meshpts,meshpts,fxn, params)))}
# Fecundity
F_k <- function(h, recruit_dist, ns) {
    seedlings <- expected_seedlings(ns)
    offspring <- recruit_dist*seedlings#*h #(no need to multiply by
        h)
    return(offspring)}
F_k_mod <- function(h,recruit_dist,ns) {
    per_capita <- per_capita_offspring(ns)*mean_seeds(1)
    matrix <- h*outer(recruit_dist, rep(per_capita, times=m),FUN="*")
    return(matrix)
}
```

```
## llll
### - Functions to update state
# For both types of simulations
update_seeds <- function(h,nr) {
    ns <- sum(nr)*h*mean_seeds(1)
    return(ns)}
update_adults <- function(h,meshpts, recruit_dist, dens,nt,nr,modl)
        {
    ns <- sum(nr)*h*mean_seeds(1)
    recruits <- F_k(h,recruit_dist, ns)
    survivors <- P_k(h,meshpts, dens,modl) %*% nt
    nt1 <- recruits + survivors
    return(nt1)}
update_adults_ws <- function(h, meshpts, recruit__dist, dens, ns, nt, nr
        ,modl) {
    #ns <- sum(nr)*h*mean_seeds (1)
        recruits <- F_k(h,recruit_dist, ns)
        survivors <- P_k(h,meshpts,dens,modl) %*% nt
        nt1 <- recruits + survivors
        return(nt1)}
update_flowering <- function(h,meshpts,dens,nt,modl) {
        nr <- R_k(h,meshpts, dens,modl) %*% nt
        return(nr)}
update_density <- function(h, weights, nt, nr, area_frac, quality) {
        if (is.na(quality) | sum(nt+nr)==0) {return(0)}
```

```
    else {return((weights %*% (nt+nr)*h)/(area_frac*quality))}}
# Function to simulate population dynamics
master_local_simulation_ws <- function(L,U,m, num_steps, ns_init,
        nt_init,nr_init, modl) {
###
## -- Initial processing
# Design mesh
h <- (U-L)/m
meshpts <- L + (1:m)*h - h/2
weights <- (meshpts/100)^mass_exp
recruit_dist <- recruit_size(meshpts)
print("Status: meshes made")
###
## -- Run the simulation
# Initialize vectors
ns <- rep (0,times=num_steps+1)
nt <- array (0,dim=c (num_steps+1,m))
nr <- array (0, dim=c (num_steps+1,m))
dt <- rep(0,times=num_steps+1)
# Initial values
ns[1] <- ns_init
nt[1,] <- nt_init
nr[1,]<- nr_init
dt[1] <- update_density(h,weights, nt[1,], nr [1,], 1,1)
print("Status:ssimulationuinitialized")
# Loop through each time step
    for (t in 1:num_steps) {
        nt[t+1,] <- update_adults_ws(h, meshpts,recruit_dist, dt[t],ns[
            t],nt[t,], nr[t,], modl)
        nr[t+1,]<- update_flowering(h,meshpts,dt[t],nt[t,], modl)
```

```
        ns[t+1] <- update_seeds(h, nr[t+1,])
        dt[t+1] <- update_density(h,weights, nt[t+1,],nr[t+1,], 1,1)
        print(ns[t+1])
    }
    # Return the density over time
    return(list(" adults"=nt," flowering"=nr," densities"=dt," seeds"=
        ns ))
```

\}

## B. 2 Diagnostics

## B.2.1 Form of Growth

```
# Library
library(lme4)
###
## -- Other forms for growth
# Functions for the forms
quad_fun <- function(x, coefs) {
    a<-coefs [1]; b<-coefs[2]; return(a*x^2+b*x)}
quad3_fun <- function(x, coefs) {
    a<-coefs[1]; b<-coefs [2]; c<-coefs[3]; return(a*x^2+b*x+c)}
rick_fun <- function(x, coefs) {
    a<-coefs [1]; b<-coefs[2]; return(a*x*exp(-b*x))}
skel_fun <- function(x, coefs) {
    a<-coefs[1]; b<-coefs [2]; return(a*(1-exp(-b*x)))}
bvht_fun <- function(x, coefs) {
    a<-coefs [1]; b<-coefs[2]; return(a*x/(b+x))}
powr_fun <- function(x, coefs) {
    a<-coefs[1]; b<-coefs [2]; return(a*x^b)}
# Fit the fits
dat_asymp <- hmdata_adult[!is.na(hmdata_adult $ Height2002),]
```

line_fit $<-\operatorname{lm}(h 1 \sim$ h, data=dat_asymp)
quad_fit <- nls (h1 ~ quad_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
quad3_fit<- lm(h1~poly (h, 2), data=dat_asymp)
rick_fit <- nls(h1~ rick_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=0.01))$
skel_fit <- nls(h1~ skel_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=100, b=0.01))$
bvht_fit <- nls(h1 ~ bvht_fun(h,c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
powr_fit $<-$ nls (h1 ~ powr_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
\# AIC for the above fits
round ( $\operatorname{AIC}\left(\right.$ line fit $\left.^{\text {fit }}\right)$ )
round (AIC(quad_fit))
round (AIC(quad3_fit))
round (AIC(rick_fit))
round (AIC(skel_fit))
round (AIC(bvht_fit))
round (AIC (powr_fit))
\# Fit some more fits
dat_asymp $<-$ hmdata_flow
line_fit <- lm(h1 ~h, data=dat_asymp)
quad_fit <- nls(h1 ~ quad_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
quad3_fit<- lm(h1~poly (h, 2), data=dat_asymp)
rick_fit <- nls(h1 ~ rick_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=0.01))$
skel_fit <- nls(h1 ~ skel_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=100, b=0.01))$
bvht_fit <- nls(h1 ~ bvht_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
powr_fit <- nls(h1 ~ powr_fun(h, c(a,b)), data=dat_asymp, start=
list $(a=1, b=1)$ )
\# AIC for even more fits

```
round(AIC(line _fit))
round(AIC(quad_fit))
round(AIC(quad3_fit))
round(AIC(rick_fit))
round(AIC(skel_fit))
round (AIC(bvht_fit))
round(AIC(powr_fit))
```


## B.2.2 Density-Dependence and Site-Specific Effects

```
###
## -- Test density-dependence and site-specific effects
# Survival
surv_fit_orig <- glm(survived ~ h, data=dat, family="binomial")
surv_fit_dens <- glm(survived ~ h+dens, data=dat, family="
        binomial")
surv_fit_site <- glmer(survived ~ h+(1|h), data=dat, family="
    binomial")
AIC(surv_fit_orig)
AIC(surv_fit_dens)
AIC(surv_fit_site)
# Reproduction
repr_fit_orig <- glm(reproduced ~ h, data=dat, family="binomial")
repr_fit_dens <- glm(reproduced ~ h+dens, data=dat, family="
    binomial")
repr_fit_site<- glmer(reproduced ~ h+(1|h), data=dat, family="
    binomial")
AIC(repr_fit_orig)
AIC(repr_fit_dens)
AIC(repr_fit_site)
lrtest(repr_fit_orig,repr_fit_dens)
```

```
# Vegetative growth
grow_fit_a_orig <- lm(h1~h, data=hmdata_adult)
grow_fit_a_dens <- lm(h1~h+dens, data=hmdata_adult)
grow_fit_a_site <- lmer(h1~h+(1|h), data=hmdata_adult)
AIC(grow_fit_a_orig)
AIC(grow_fit_a_dens)
AIC(grow_fit_a_site)
lrtest(grow_fit_a_orig,grow_fit_a_dens)
# Vegetative growth
grow_fit_r_orig <- lm(h1~h, data=hmdata_flow)
grow_fit_r_dens <- lm(h1~h+dens, data=hmdata_flow)
grow_fit_r_site <- lmer(h1~h+(1|h), data=hmdata_flow)
AIC(grow_fit_r_orig)
AIC(grow_fit_r_dens)
AIC(grow_fit_r_site)
lrtest(grow_fit_r_orig, grow_fit_r_dens)
lrtest(grow_fit_r_orig,grow_fit_r_site)
```


## B.2.3 Normality of Residuals and Homoskedasticity

```
###
## - Normality of residuals and heteroskedasticity
# For vegetative plants
plot(hmdata_adult $h,grow_fit_a$residuals)
shapiro.test(grow_fit_a$residuals)
bptest(grow_fit_a)
# For reproductive plants
plot (hmdata_flow $h,hmdata_flow $h1)
plot(hmdata_flow $h, grow_fit_r$residuals)
shapiro.test(grow_fit_r$residuals)
bptest(grow_fit_r)
```


## B.2.4 Log-normality of the Recruit Size Distribution

```
###
## - Test recruit fit
shapiro.test(log(seedlingdata$ Height2003))
sd(log(seedlingdata$Height2003))
qqnorm(log(seedlingdata$Height2003))
```


## B. 3 Bootstrapping

## B.3.1 Bootstrap population growth rate

```
# Library to test if intervals overlap
library (DescTools)
# Function to simulate population given a bootstrap replicate
bootstrap_helper <- function(adult_dat, seedling_dat,modl,L,U,m) {
###
## -- The data set to perform regression on
# Which subset of data to use
dat <- adult_dat
# Use this one for growth
dat_adult <- dat[dat$Stage2003!=4 & dat $Stage2003!=5,]
dat_flow <- dat[dat$Stage2003==4,]
###
## - Preamble to finding equilibria
# Specify mesh and weights
h <- (U-L)/m; meshpts <- L + (1:m)*h - h/2
weights <- (meshpts/100)^mass_exp
```

```
###
## -- Survival and reproduction
# Survival
surv_fit <- glm(survived ~h, data=dat, family="binomial")
p_s <- function(h) {return(predict(surv_fit, data.frame(h=h),
    type="response"))}
# Growth
if (modl="rl") {
    repr_fit <- glm(reproduced ~h, data=dat, family="binomial")
    p_r <- function(h, dens) {
        return(predict(repr_fit, data.frame(h=h), type="response"))
            }
} else if (modl=" ac") {
    repr_fit <- glm(reproduced ~h+dens, data=dat, family="binomial
            ")
    p_r<- function(h, dens) {
        return(predict(repr_fit, data.frame(h=h,dens=dens), type="
            response"))}
}
###
## -- Growth
# Fit growth
grow_fit_a <- lm(h1~h, data=dat_adult)
grow_fit_r <- lm(h1 ~h, data=dat_flow)
# Growth kernel
g_a <- function(h1,h) {
    h1bar <- predict(grow_fit_a, newdata=data.frame(h=h), type="
        response")
    sdhat <- sqrt(sum(residuals(grow_fit_a)^2)/df.residual(grow_
        fit_a))
    return(dnorm(h1, mean=h1bar, sd=sdhat))}
g_r <- function(h1,h) {
```

55 h1bar $<-$ predict (grow_fit_r, newdata=data.frame(h=h), type=" response")

```
    sdhat <- sqrt(sum(residuals(grow_fit_r)^2)/df.residual(grow_
```

        fit_r))
    return (dnorm(h1, mean=h1bar, \(s d=\) sdhat) \()\}\)
    ```
###
## -- Number of recruits
# Seedling survival parameters
# max_seedlings <- 127.5 #max_seedlings <- mean(seedling_pos$
    Seedlings2003)
# prob_surv <- 0.32
alpha <- 127.5
beta <- 273.4375
# Function to return expected number of seedlings given number
        of flowering plants
expected_seedlings <- function(ns) {
    num_seedlings <- alpha*ns/(beta+ns)
    return(num_seedlings)
}
per_capita_offspring <- function(ns) {
    num_offspring <- alpha/(beta+ns)
    return(num_offspring)
}
# Fit recruit size to a log-normal distribution
r_size <- fitdistr(seedling_dat$Height2003, "lognormal")$
        estimate
# Function for probability of recruit size
recruit_size <- function(zvec) {return(dlnorm(zvec, meanlog=r_
        size["meanlog"], sdlog=r_size["sdlog"], log=F))}
###
## - Find kernels
```

\# Survival-growth
P_k <- function (h, meshpts, dens) \{
fxn $<-$ function (z1,z, dens) \{
return $\left.\left(p_{-}(z) *\left(1-p_{-}(z, d e n s)\right) * g_{-}(z 1, z)\right)\right\}$
return (h* (outer (meshpts, meshpts,fxn, dens)) ) \}
\# Reproduction
R_k <- function (h, meshpts, dens) \{
fxn $<-$ function (z1, z, dens) \{ return (p_s (z)*p_r (z, dens) *g_r (z1, z) ) \}
return (h* (outer (meshpts, meshpts, fxn, dens)) ) \}
\# Fecundity
F_k <- function (h, recruit_dist, ns) \{
seedlings $<-$ expected_seedlings (ns)
offspring <- recruit_dist*seedlings\#*h \#(no need to multiply by h?)
return(offspring) \}
\#\#\#
\#\# - Functions to update state
update_seeds <- function (h, nr) \{
ns $<-\operatorname{sum}(\mathrm{nr}) *$ h $*$ mean_seeds (1)
return(ns) $\}$
update_adults_ws <- function (h, meshpts, recruit_dist, dens, ns, nt, nr, modl) \{
recruits <- F_k(h, recruit_dist, ns)
survivors <- P_k(h, meshpts, dens) \%*\% nt
nt1 <- recruits + survivors
return(nt1) \}
update_flowering $<-$ function (h, meshpts, dens, nt, modl) \{
nr <- R_k(h, meshpts, dens) \%*\% nt
return(nr) \}

```
update_density <- function(h, weights, nt,nr, area_frac, quality)
        {
    if (is.na(quality) | sum(nt+nr)==0) {return(0)}
    else {return((weights %*% (nt+nr)*h)/(area_frac*quality))}}
###
## -- Function to iterate
master_local_simulation_ws <- function(L,U,m, num_steps, ns_
        init,nt_init,nr_init, modl) {
    # Design mesh
    h <- (U-L)/m
    meshpts <- L + (1:m)*h - h/2
    weights <- (meshpts/100)^mass_exp
    recruit_dist <- recruit_size(meshpts)
    print("Status:^meshesumade")
    # Initialize vectors
    ns <- rep(0,times=num_steps+1)
    nt <- array (0,dim=c(num_steps+1,m))
    nr <- array (0, dim=c (num_steps+1,m))
    dt <- rep (0,times=num_steps+1)
    # Initial values
    ns[1] <- ns_init
    nt[1,]<- nt_init
    nr[1,]<- nr_init
    dt[1] <- update_density(h,weights,nt[1,],nr[1,], 1,1)
    # Loop through each time step
    for (t in 1:num_steps) {
        nt[t+1,]<- update_adults_ws(h,meshpts, recruit_dist, dt[t],
            ns[t],nt[t,],nr[t,], modl)
        nr[t+1,]<- update_flowering(h,meshpts,dt[t], nt[t,], modl)
        ns[t+1] <- update_seeds(h, nr [t+1,])
        dt[t+1] <- update_density(h, weights, nt [t+1,],nr[t+1,],
            1,1)
```

```
        }
        # Return the density over time
        return(list("adults"=nt,"flowering"=nr,"densities"=dt,"seeds"
            =ns))
    }
    ###
    ## - Run the simulation
    # Initialize
    num_steps <- 30
    ns_init <- 1
    nt_init <- rep(0,length=m)
    nr_init <- rep(0,length=m)
    # Run simulations
    sim_res <- master_local_simulation_ws(L,U,m, num_steps, ns_init
        ,nt_init,nr_init, modl)
    ###
    ## - Return things
    # Return the number of seeds and distribution of adult and
        flowering plants
    return(list("seeds"=sim_res$seeds," adults"=sim_res$adults,"
        flowering"=sim_res$flowering))
}
###
## - Function for bootstrapping
bootstrap_main <- function(dat, N, modl) {
    # Initialize things
    L<- 5; U <- 550; m <- 50
    h <- (U-L)/m; meshpts <- L + (1:m)*h - h/2
    weights <- (meshpts/100)^mass_exp
```

```
# Number of rows to sample from
nr <- nrow(dat)
# Vectors to store initial seed values
seeds0 <- array (NA, dim=c (N))
seeds1<- array (NA, dim=c (N))
seeds2 <- array (NA, dim=c(N))
seeds3 <- array (NA, dim=c(N))
seeds4<- array (NA, dim=c(N))
seeds5 <- array (NA, dim=c (N))
seeds6 <- array (NA, dim=c(N))
seeds7<- array (NA, dim=c(N))
# Vectors to store asymptotic seed values
seeds29 <- array (NA, dim=c (N))
seeds30<- array (NA, dim=c(N))
# Vectors to store heights of vegetative plants, reproductive
    plants, and biomass
# NOTE: 's' means 'simulated' while 'o' means 'observed'
height_as29<- array (NA, dim=c(N))
height_fs29<- array (NA, dim=c(N))
height_as30 <- array (NA, dim=c(N))
height_fs30<- array (NA, dim=c(N))
height_ao <- array (NA, dim=c(N))
height_fo <- array (NA, dim=c(N))
biomass <- array (NA, dim=c(N))
# Index
i = 1
# Bootstrap step
while (i <= N) {
    # Randomly sample from data set
    indices <- sample.int(nr, nr, replace=T)
    dat_new <- dat[indices,]
    dat_adult <- dat_new[dat_new$Stage2003!=4 & dat_new$Stage2003
        !=5,]
    dat_flow <- dat_new[dat_new$Stage2003==4,]
    # Do bootstrapping
    out <- bootstrap_helper(dat_new[dat_new$class!="Seedling",],
        dat_new[dat_new$class=" Seedling",],modl,L,U,m)
    # Give useful names to final versions
```

```
    num_steps <- length(out$seeds)
    adlt_2last <- out$adults[num_steps - 1,]/sum(out$adults[num_
        steps-1,])
    flow_2last <- out$flowering[num_steps - 1,]/sum(out$ flowering[
        num_steps - 1,])
    adlt_last <- out$adults[num_steps,]/sum(out$adults[num_steps
        ,])
    flow_last <- out$flowering[num_steps,]/sum(out $ flowering[num
        _steps,])
    # Get statistics
    seeds0[i] <- out$seeds[1]
    seeds1[i] <- out$seeds[2]
    seeds2[i] <- out$seeds[3]
    seeds3[i] <- out$seeds[4]
    seeds4[i] <- out$seeds[5]
    seeds5[i] <- out$seeds[6]
    seeds6[i] <- out$seeds[7]
    seeds7[i] <- out$seeds[8]
    seeds29[i] <- out$seeds[num_steps - 1]
    seeds30[i] <- out$seeds[num_steps]
    biomass[i] <- weights %**% (out$adults[num_steps,]+out$
        flowering[num_steps,]) * h
    height_as29[i] <- meshpts %*% adlt_2last
    height_fs29[i] <- meshpts %*% flow_2last
    height_as30[i] <- meshpts %*% adlt_last
    height_fs30[i] <- meshpts %*% flow_last
    height_ao[i] <- mean(dat_adult$h1, na.rm=T)
    height_fo[i] <- mean(dat_flow$h1)
    print(i)
    i}<- i+
}
return(list("seeds0"=seeds0,"seeds1"=seeds1,"seeds2"=seeds2 ,"
    seeds3"=seeds3,
                "seeds4"=seeds4,"seeds5"=seeds5,"seeds6"=seeds6,"
                seeds7"=seeds7,
            "seeds29"=seeds29,"seeds30"=seeds30,"biomass"=
                biomass,
```

| 255 | "height_as29"=height_as29, "height_fs 29 " $=$ height_fs 29 |
| :---: | :---: |
| 256 | $" \text { height_as } 30 "=\text { height_as30, "height_fs } 30 "=h e i g h t \_f s 30$ |
| 257 | "height_ao"=height_ao, "height_fo"=height_fo)) |
| 258 | \} |
| 259 |  |
| 260 |  |
| 261 | \#\#\# |
| 262 | \#\# - Run bootstrap simulations |
| 263 |  |
| 264 | \# Subset data into high and low |
| 265 | dat $<-$ rbind (hmdata, seedlingdata) |
| 266 |  |
| 267 | \# Compute the bootstrap confidence intervals |
| 268 | \# Took 65 seconds when $\mathrm{N}=100$ for each |
| 269 | tm <- proc.time () |
| 270 | boot_distn_rl $<-$ bootstrap_main(dat, 500, , rl") |
| 271 | boot_distn_ac <- bootstrap_main (dat, 500, "ac") |
| 272 | print (proc.time ()-tm) |
| 273 |  |
| 274 |  |
| 275 | \#\#\#\# |
| 276 | \#\# - Seed informations |
| 277 |  |
| 278 | \# Get seed numbers |
| 279 | s0_distn_rl <- boot_distn_rl\$seeds0 |
| 280 | s1_distn_rl <- boot_distn_rl\$seeds1 |
| 281 | s2_distn_rl <- boot_distn_rl\$seeds2 |
| 282 | s3_distn_rl <- boot_distn_rl\$seeds 3 |
| 283 | s4_distn_rl <- boot_distn_rl\$seeds 4 |
| 284 | s5_distn_rl <- boot_distn_rl\$seeds5 |
| 285 | s6_distn_rl <- boot_distn_rl\$seeds6 |
| 286 | s7_distn_rl <- boot_distn_rl\$seeds7 |
| 287 | s29_distn_rl <- boot_distn_rl\$seeds29 |
| 288 | s30_distn_rl $<-$ boot_distn_rl\$seeds 30 |
| 289 |  |
| 290 | s0_distn_ac $<-$ boot_distn_ac\$seeds0 |


| 291 | s1_distn_ac <- boot_distn_ac\$seeds1 |
| :---: | :---: |
| 292 | s2_distn_ac <- boot_distn_ac\$seeds2 |
| 293 | s3_distn_ac <- boot_distn_ac\$seeds 3 |
| 294 | s4_distn_ac <- boot_distn_ac\$seeds4 |
| 295 | s5_distn_ac <- boot_distn_ac\$seeds5 |
| 296 | s6_distn_ac <- boot_distn_ac\$seeds6 |
| 297 | s7_distn_ac <- boot_distn_ac\$seeds7 |
| 298 | s29_distn_ac <- boot_distn_ac\$seeds 29 |
| 299 | s30_distn_ac <- boot_distn_ac \$seeds 30 |
| 300 |  |
| 301 | \# Bootstrap means |
| 302 | s0_boot_mu_rl <- mean(s0_distn_rl) |
| 303 | s1_boot_mu_rl $<-$ mean(s1_distn_rl) |
| 304 | s2_boot_mu_rl <- mean(s2_distn_rl) |
| 305 | s3_boot_mu_rl <- mean(s3_distn_rl) |
| 306 | s4_boot_mu_rl <- mean(s4_distn_rl) |
| 307 | s5_boot_mu_rl <- mean(s5_distn_rl) |
| 308 | s6_boot_mu_rl <- mean(s6_distn_rl) |
| 309 | s7_boot_mu_rl <- mean(s7_distn_rl) |
| 310 | s29_boot_mu_rl $<-\operatorname{mean}(\mathrm{s} 29 \ldots$ distn_rl) |
| 311 | s30_boot_mu_rl $<-\operatorname{mean}\left(\mathrm{s} 30 \_\right.$distn_rl) |
| 312 |  |
| 313 | s0_boot_mu_ac <- mean(s0_distn_ac) |
| 314 | s1_boot_mu_ac <- mean(s1_distn_ac) |
| 315 | s2_boot_mu_ac <- mean(s2_distn_ac) |
| 316 | s3_boot_mu_ac <- mean(s3_distn_ac) |
| 317 | s4_boot_mu_ac <- mean(s4_distn_ac) |
| 318 | s5_boot_mu_ac <- mean(s5_distn_ac) |
| 319 | s6_boot_mu_ac <- mean(s6_distn_ac) |
| 320 | s7_boot_mu_ac <- mean(s7_distn_ac) |
| 321 | s29_boot_mu_ac <- mean (s29_distn_ac) |
| 322 | s30_boot_mu_ac <- mean (s30_distn_ac) |
| 323 |  |
| 324 | \# Bootstrap confidence intervals |
| 325 | s0_boot_ci_rl <- quantile(s0_distn_rl, probs $=$ c (0.025, 0.975$)$ ) |
| 326 | s1_boot_ci_rl <- quantile(s1_distn_rl, probs=c (0.025, 0.975$)$ ) |
| 327 | s2_boot_ci_rl <- quantile (s2_distn_rl, probs=c (0.025, 0.975) |
| 328 | s3_boot_ci_rl <- quantile(s3_distn_rl, probs=c (0.025,0.975) |


| 329 |  |
| :---: | :---: |
| 330 |  |
| 331 |  |
| 332 | s7_boot_ci_rl <- quantile(s7_distn_rl, probs=c (0.025, 0.975$)$ ) |
| 333 | s29_boot_ci_rl $<-$ quantile (s29_distn_rl, probs $=\mathrm{c}(0.025,0.975)$ ) |
| 334 |  |
| 335 |  |
| 336 | s0_boot_ci_ac <- quantile (s0_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 337 | s1_boot_ci_ac <- quantile(s1_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 338 | s2_boot_ci_ac $<-$ quantile (s2_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 339 | s3_boot_ci_ac $<-$ quantile (s3_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 340 | s4_boot_ci_ac $<-$ quantile (s $4_{\text {_ distn_ac, }}$ probs $=\mathrm{c}(0.025,0.975)$ ) |
| 341 | s5_boot_ci_ac $<-$ quantile $^{\text {a }} 5$ _ distn_ac, probs=c $(0.025,0.975)$ ) |
| 342 |  |
| 343 | s7_boot_ci_ac <- quantile (s7_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 344 | s29_boot_ci_ac <- quantile (s29_distn_ac, probs=c (0.025, 0.975$)$ ) |
| 345 |  |
| 346 |  |
| 347 |  |
| 348 | \#\#\# |
| 349 | \#\# - Seed table |
| 350 |  |
| 351 | \# Compare second year |
| 352 | s2_boot_mu_rl |
| 353 | s2_boot_ci_rl |
| 354 |  |
| 355 | s2 _ boot_mu_ac |
| 356 | s2_boot_ci_ac |
| 357 |  |
| 358 | \# Compare third year |
| 359 | s3_boot_mu_rl |
| 360 | s3_boot_ci_rl |
| 361 |  |
| 362 | s3_boot_mu_ac |
| 363 | s3_boot_ci_ac |
| 364 |  |
| 365 | \# Compare fourth year |
| 366 | s4_boot_mu_rl |


| 367 | s4_boot_ci_rl |
| :---: | :---: |
| 368 |  |
| 369 | s4_boot_mu_ac |
| 370 | s4_boot_ci_ac |
| 371 |  |
| 372 | \# Compare fifth year |
| 373 | s5_boot_mu_rl |
| 374 | s5_boot_ci_rl |
| 375 |  |
| 376 | s5_boot_mu_ac |
| 377 | s5_boot_ci_ac |
| 378 |  |
| 379 | \# Compare sixth year |
| 380 | s6_boot_mu_rl |
| 381 | s6_boot_ci_rl |
| 382 |  |
| 383 | s6_boot_mu_ac |
| 384 | s6_boot_ci_ac |
| 385 |  |
| 386 | \# Compare seventh year |
| 387 | s7_boot_mu_rl |
| 388 | s7_boot_ci_rl |
| 389 |  |
| 390 | s7_boot_mu_ac |
| 391 | s7_boot_ci_ac |
| 392 |  |
| 393 | \# Compare twenty-ninth year |
| 394 | s29_boot_mu_rl |
| 395 | s29_boot_ci_rl |
| 396 |  |
| 397 | s29_boot_mu_ac |
| 398 | s29_boot_ci_ac |
| 399 |  |
| 400 | \# Compare thirtieth year |
| 401 | s30_boot_mu_rl |
| 402 | s30_boot_ci_rl |
| 403 |  |
| 404 | s30_boot_mu_ac |


| 405 | s30_boot_ci_ac |
| :---: | :---: |
| 406 |  |
| 407 | \#\#\# |
| 408 | \#\# - Check if the differences are significant |
| 409 |  |
| 410 | s2_boot_ci_rl \%overlaps\% s2_boot_ci_ac |
| 411 | s3_boot_ci_rl \%overlaps\% s3_boot_ci_ac |
| 412 | s4_boot_ci_rl \%overlaps\% s4_boot_ci_ac |
| 413 | s5_boot_ci_rl \%overlaps\% s5_boot_ci_ac |
| 414 | s6_boot_ci_rl \%overlaps\% s6_boot_ci_ac |
| 415 | s7_boot_ci_rl \%overlaps\% s7_boot_ci_ac |
| 416 |  |
| 417 | s29_boot_ci_rl \%overlaps\% s29_boot_ci_ac |
| 418 | s30_boot_ci_rl \%overlaps\% s30_boot_ci_ac |
| 419 |  |
| 420 |  |
| 421 | \#\#\# |
| 422 | \#\# - Height information |
| 423 |  |
| 424 | \# Get heights and biomass |
| 425 | has29_distn_rl<- boot_distn_rl\$height_as 29 |
| 426 | has29_distn_ac <- boot_distn_ac \$height_as 29 |
| 427 | hfs 29 _ distn_rl $<-$ boot_distn_rl\$height_fs 29 |
| 428 |  |
| 429 |  |
| 430 | has30_distn_rl $<-$ boot_distn_rl \$height_ as $30^{\text {d }}$ |
| 431 | has30_distn_ac <- boot_distn_ac \$height_as 30 |
| 432 |  |
| 433 |  |
| 434 |  |
| 435 | hao_distn_rl $<-$ boot_distn_rl\$height_ao |
| 436 |  |
| 437 | hfo_distn_rl $<-$ boot_distn_rl\$height_fo |
| 438 |  |
| 439 |  |
| 440 | b_distn_rl $^{\text {c- }}$ boot_distn_rl\$biomass |
| 441 | $b_{-}$distn_ac $<-$boot_distn_ac \$ biomass |
| 442 |  |

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```
# Bootstrap means
has29_boot_mu_rl <- mean(has29_distn_rl)
has29_boot_mu_ac <- mean(has29_distn_ac)
hfs29_boot_mu_rl<- mean(hfs29_distn_rl)
hfs29_boot_mu_ac<< mean(hfs29_distn_ac)
has30_boot_mu_rl<- mean(has30_distn_rl)
has30_boot_mu_ac<- mean(has30_distn_ac)
hfs30_boot_mu_rl<< mean(hfs30_distn_rl)
hfs30_boot_mu_ac<< mean(hfs30_distn_ac)
hao_boot_mu_rl <- mean(hao_distn_rl)
hao_boot_mu_ac <- mean(hao_distn_ac)
hfo_boot_mu_rl<- mean(hfo_distn_rl)
hfo_boot_mu_ac <- mean(hfo_distn_ac)
b_boot_mu_rl <- mean(b_distn_rl)
b_boot_mu_ac <- mean(b_distn_ac)
# Bootstrap confidence intervals
has29_boot_ci_rl<- quantile(has29_distn_rl, probs=c(0.025,0.975))
has29_boot_ci_ac<- quantile(has29_distn_ac, probs=c (0.025,0.975))
hfs29_boot_ci_rl<- quantile(hfs29_distn_rl, probs=c (0.025,0.975))
hfs29_boot_ci_ac<< quantile(hfs29_distn_ac, probs=c (0.025,0.975))
has30_boot_ci_rl<- quantile(has30_distn_rl, probs=c (0.025,0.975))
has30_boot_ci_ac<- quantile(has30_distn_ac, probs=c (0.025,0.975))
hfs30_boot_ci_rl<- quantile(hfs30_distn_rl, probs=c(0.025,0.975))
hfs30_boot_ci_ac<< quantile(hfs30_distn_ac, probs=c (0.025,0.975))
hao_boot_ci_rl<- quantile(hao_distn_rl, probs=c(0.025,0.975))
hao_boot_ci_ac<- quantile(hao_distn_ac, probs=c(0.025,0.975))
hfo_boot_ci_rl<- quantile(hfo_distn_rl, probs=c(0.025,0.975))
hfo_boot_ci_ac<- quantile(hfo_distn_ac, probs=c(0.025,0.975))
b_boot_ci_rl<- quantile(b_distn_rl, probs=c(0.025,0.975))
b_boot_ci_ac<- quantile(b_distn_ac, probs=c(0.025,0.975))
```


## B. 4 Determine Equilibrium and Stablility

## B.4.1 Determine equilibrium

1

```
library(grDevices)
###
## - Find dominant eigenvalue of P
L<- 5; U <- 550; m <- 1000
h <- (U-L)/m; meshpts <- L + (1:m)*h - h/2
P<- P_k(h,meshpts,1000000," ac" )
Re(eigen(P,only.values=T)$values[1])
###
## - Code discretization
# Params
L<- 5; U <- 550; m <- U-L
h <- (U-L)/m; meshpts <- L + (1:m)*h - h/2
weights <- (meshpts/100)^mass_exp
recruit_dist <- recruit_size(meshpts)
# Which model to use
modl <- "rl"
###
## -- Various helper functions
# Helper function for A(b)
A <- function(b,modl) {
    P_mat <- P_k(h,meshpts,b,modl)
    R_mat <- R_k(h,meshpts,b,modl)
    Z_mat <- matrix (0, nrow=m, ncol=m)
```

```
    tmp <- cbind(rbind(P_mat,R_mat),rbind(Z_mat,Z_mat))
    return(tmp)
}
# Helper function for n_s^*
find_ns <- function(pe) {
    ns_star <- 0
    if (pe<alpha) {
        ns_star <- (alpha-pe*beta)/pe}
        #ns_star <- pe*beta/(alpha-pe)}
    return(ns_star)
}
###
## -- Find equilibium for recruit-limited model
# Stability radius
q<- t(rbind(matrix (0, nrow=m, ncol=1),matrix (h*mean_seeds(1) , nrow=
    m, ncol=1)))
c0 <- rbind(as.matrix(recruit_size(meshpts)),matrix(0,nrow=m, ncol
    =1))
tmp <- A(0," rl")
fund_mat <- solve(diag (2*m)-tmp)
pe_rl <- 1/(q%*%fund_mat%*%c0)
# Equilibrium seeds
ns_eq_rl<- find_ns(pe_rl)
# Equilibrium
n_eq_rl <- drop(pe_rl*ns_eq_rl) * (fund_mat%*%c0)
nt_eq_rl<- n_eq_rl[1:m]
nr_eq_rl <- n_eq_rl[(m+1):(2*m)]
# Equilibrium biomass
b_eq_rl <- (nt_eq_rl+nr_eq_rl)%**%weights
b_eq_rl
```

```
###
## - Find equilibrium dens as a function of $b$
# Initialize
N <- 100
max_dens <- 4000
b_vec <- as.matrix(seq(1,max_dens, length=N))
pe_vec <- array (0, dim=N)
ns_vec <- array (0, dim=N)
n_eq_vec <- array (0,dim=c (N, 2*m))
bs_new <- array (0, dim=N)
# Precompute these vectors for speeeeed
q<- t(rbind(matrix(0, nrow=m, ncol=1),matrix (h*mean_seeds(1) , nrow=
    m, ncol=1)))
c0 <- rbind(as.matrix(recruit_size(meshpts)),matrix(0,nrow=m, ncol
    =1))
# p^*_e = 1/(q^T (I-A(b) )^{-1} c_0)
ptm <- proc.time()
for (i in 1:N) {
    b <- b_vec[i]
    tmp <- A(b," ac")
    fund_mat <- solve(diag (2*m)-tmp)
    pe_vec[i] <- 1/(q%*%fund_mat%*%c0)
    ns_vec[i] <- find_ns(pe_vec[i])
    n_eq_vec[i,] <- fund_mat%*%c0 *(pe_vec[i]*ns_vec[i])
    bs_new[i] <- update_density(h, weights, n_eq_vec [i, 1:m],n_eq_vec[
        i,(m+1):(2*m)],1,1)
    print(i)
}
print(proc.time()-ptm)
###
## -- Find the biomass which gives the equilibrium size
    distribution
```

```
# Params
L}<-5; U<- 550; m <- U-L
h}<-(\textrm{U}-\textrm{L})/\textrm{m}; meshpts <- L + (1:m)*h - h/2
weights <- (meshpts/100)^ mass_exp
recruit_dist <- recruit_size(meshpts)
# Initialize
b_l <- 2142
b_u <- 2146
b_m<- (b_l+b_u)/2
b_approx <- 2500
err<< b_approx-b_m
tol <- 0.00001
# Precompute these vectors for speeeeed
q<- t(rbind(matrix (0, nrow=m, ncol=1), matrix (h*mean_seeds (1) , nrow=
    m,ncol=1)))
c0<- rbind(as.matrix(recruit_size(meshpts)), matrix (0, nrow=m, ncol
    =1))
# Run the loop
while (abs(err) > tol) {
    if (err>0) {b_l<- b_m}
    if (err<0) {b_u<< b_m}
    b_m<- (b_l+b_u)/2
    fund_mat <- solve(diag (2*m)-A(b_m," ac"))
    pe<-1/(q%*%fund_mat%*%c0)
    ns <- find_ns(pe)
    bar_n <- fund_mat%*%c0 * drop (pe*ns)
    b_approx <- update_density (h, weights, bar_n[1:m], bar_n[(m+1):(2*
            m)],1,1)
    err<- b_approx-b_m
    print(err)
}
b _ approx
# Give better names to the final values
pe_ac <- pe
```

```
ns_eq_ac <- ns
n_eq_ac <- bar_n
nt_eq_ac <- n_eq_ac[1:m]
nr_eq_ac <- n_eq_ac [(m+1):(2*m)]
b_eq_ac <- b_approx
###
## -- Save equilibria for later use in stability analysis
write.csv(n_eq_rl, file="equilibrium_rl.csv")
write.csv(n_eq_ac, file="equilibrium_ac.csv")
###
###
###
    \square
###
## -- Compare biomasses
# Equilibrium biomasses
b_eq_rl
b_eq-ac
# vs max observed biomass
max(hmdata$dens,hmdata$dens1)
###
## - Compare seed production and number of flowering plants
# Equilibrium seeds
ns_eq_rl
```

```
ns_eq_ac
# Equilibrium flowering plants
ns_eq_rl/(mean_seeds (1))
ns_eq_ac/(mean_seeds(1))
# Equilibrium flowering plants (alt...should equal above)
sum(nr_eq_rl)*h
sum(nr_eq_ac)}*
# Observed flowering plants
total_flowering <- sum(hmdata$reproduced, na.rm=T)
total_plots <- length(levels(as.factor(hmdata$plot)))
total_flowering/total_plots
###
## - Compare number of vegetative plants
# Equilibrium vegetative plants
sum(nt_eq_rl)*h
sum(nt_eq_ac)}*\textrm{h
# Observed vegetative plants (including seedlings)
hmdata_veg <- rbind(seedlingdata,hmdata_adult)
total_vegetative <- nrow(hmdata_veg)
total_vegetative/total_plots
###
## - Compare mean heights
# Mean heights from recruit-limited model
sum(nt_eq_rl*meshpts)/sum(nt_eq_rl)
sum(nr_eq_rl*meshpts)/sum(nr_eq_rl)
# Mean heights from adult competition model
```

```
sum(nt_eq_ac*meshpts)/sum(nt_eq_ac)
sum(nr_eq_ac*meshpts)/sum(nr_eq_ac)
# Observed mean heights in 2003
mean(hmdata_veg$h1)
mean(hmdata_flow $h1)
```


## B.4.2 Determine stability of equilibrium

```
\# Helper functions for computing the Jacobian
dP_fun <- function (z1, z, dens) \{
    gamma \(<-\) as.numeric (coefficients (repr_fit_ac) [3])
    return (-gamma*p_s \((z) * p_{-} r(z, d e n s, " a c ") *\left(1-p_{-} r(z, d e n s, " a c ")\right) * g \_a(\)
        z1, z) ) \}
dR_fun <- function(z1,z, dens) \{
    gamma \(<-\) as.numeric (coefficients (repr_fit_ac) [3])
    return (gamma*p_s (z)*p_r(z, dens,"ac")*(1-p_r(z,dens,"ac"))*g_r(
        z1, z) ) \}
dF_fun<- function (z1, z, ns) \{
    return (-alpha*mean_seeds \((1) *\) recruit_size (z1)/((beta+ns)^2)) \}
\# Compute K
P_mat <- P_k(h, meshpts, b_eq_ac," ac")
R_mat <- R_k(h, meshpts, b_eq_ac," ac")
F_mat <- F_k_mod (h, recruit_dist, ns_eq_ac)
Z_mat <- matrix (0, nrow=m, ncol=m)
\(\mathrm{K}<-\operatorname{cbind}\left(\mathrm{rbind}\left(\mathrm{P} \_\right.\right.\)mat,R_mat) , rbind (F_mat, Z_mat))
\# Compute Q1
dP_mat <- h*outer (meshpts, meshpts, dP_fun, b_eq_ac)
dR_mat <- h*outer (meshpts, meshpts, dR_fun, b_eq_ac)
dZ_mat \(<-\) matrix \((0\), nrow \(=m\), ncol \(=m)\)
dK_mat <- cbind (rbind (dP_mat, dR_mat), rbind (dZ_mat, dZ_mat))
Q1 <- dK_mat \%*\% n_eq_ac
\# Compute W1
```

```
26 W1 <- as.matrix((meshpts/100)^mass_exp)
W1<- rbind(W1,W1)
# Compute Q2
dZ_mat <- matrix(0,nrow=m, ncol=m)
dF_mat <- h*outer(meshpts,meshpts,dF_fun , ns_eq_ac)
dK_mat <- cbind(rbind(dZ_mat,dZ_mat),rbind(dF_mat,dZ_mat))
Q2 <- dK_mat %*% n_eq_ac
# Compute W2
Z_vec <- matrix(0, nrow=m, ncol=1)
one_vec <- matrix(1,nrow=m, ncol=1)
W2 <- rbind(Z_vec,one_vec)*mean_seeds (1)
# Compute J
J <- K + outer(t(Q1),W1,FUN="*")[,, ,] + outer(t(Q2),W2,FUN="*")
    [, , ,]
# Determine stability
Re(eigen(J,only.values=T)$values [1])
```


## B. 5 Methods Plots

## B.5.1 Plot vital rate functions

```
###
## -- Plot vital rates
# Specify data set to use
dat <- hmdata
# Real data to plot (ticks or points as necessary)
dat_surv <- dat[!is.na(dat$survived),]
dat_repr <- dat[!is.na(dat$reproduced) ,]
dat_veg <- dat[dat$Stage2003!=4 & dat$Stage2003!=5 & ! is.na(dat$
    h) ,]
```

```
dat_flo <- dat[dat$Stage2003==4,]
# Make points that will be plotted
xmax <- 220
x_vec <- seq(0,xmax, length=100)
s_vec <- p_s(x_vec)
r_rl_vec <- p_r(x_vec,0,"rl")
r_ac_vec<- p_r(x_vec,0,"ac")
# These values are used to create segments in the growth plot
ga_min_x <- min(dat_veg$h)
ga_min_y <- as.numeric(predict(grow_fit_a,data.frame(h=ga_min_x))
    )
ga_max_x <- max(dat_veg$h)
ga_max_y <- as.numeric(predict(grow_fit_a,data.frame(h=ga_max_x))
    )
gr_min_x <- min(dat_flo$h)
gr_min_y <- as.numeric(predict(grow_fit_r, data.frame(h=gr_min_x))
    )
gr_max_x <- max(dat_flo$h)
gr_max_y <- as.numeric(predict(grow_fit_r, data.frame(h=gr_max_x))
    )
# PLOTS
par (mfrow=c (2,2))
dx <- 0.03
dy <- 1
line_frac <- 1
scale_text <- 1.3
plot(x_vec,s_vec, type="l",lwd=2, xlim=c(0,xmax), ylim=c (0,1),
        xlab=" Heightuofぃplantь(cm)", ylab="Probability", cex.lab=scale
        -text,cex.axis=scale_text)
points(dat_surv$h, dat_surv$survived, pch="|", cex=line_frac)
mtext("(a)", side=3,adj=dx, padj=dy, line= - 1.3)
plot(dat_veg$h, dat_veg$h1, pch=16, xlim=c(0,xmax),ylim=c(0,400),
```

```
        xlab=" Heightuin_2002ヶ(cm)", ylab=" Heightuin\lrcorner2003^(cm)", cex.
```

        xlab=" Heightuin_2002ヶ(cm)", ylab=" Heightuin\lrcorner2003^(cm)", cex.
        lab=scale_text, cex.axis=scale_text)
        lab=scale_text, cex.axis=scale_text)
    points(dat_flo$h, dat_flo$h1, pch=23)
points(dat_flo$h, dat_flo$h1, pch=23)
segments(ga_min_x,y0=ga_min_y, ga_max_x,y1=ga_max_y, lwd=2)
segments(ga_min_x,y0=ga_min_y, ga_max_x,y1=ga_max_y, lwd=2)
segments(gr_min_x,y0=gr_min_y, gr_max_x,yl=gr_max_y, lwd=2)
segments(gr_min_x,y0=gr_min_y, gr_max_x,yl=gr_max_y, lwd=2)
mtext("(b)", side=3,adj=dx, padj=dy, line= -1.3)
mtext("(b)", side=3,adj=dx, padj=dy, line= -1.3)
plot(x_vec,r_rl_vec, type="l", lwd=2, xlim=c (0,xmax), ylim=c (0, 1),
plot(x_vec,r_rl_vec, type="l", lwd=2, xlim=c (0,xmax), ylim=c (0, 1),
xlab="Heightヶofぃplantь(cm)", ylab="Probability", cex.lab=scale
xlab="Heightヶofぃplantь(cm)", ylab="Probability", cex.lab=scale
text, cex.axis=scale_text)
text, cex.axis=scale_text)
points(dat_repr$h,dat_repr$reproduced, pch="|", cex=line_frac)
points(dat_repr$h,dat_repr$reproduced, pch="|", cex=line_frac)
mtext("(c)", side=3,adj=dx, padj=dy, line= -1.3)
mtext("(c)", side=3,adj=dx, padj=dy, line= -1.3)
plot(x_vec,r_ac_vec, type="l", lwd=2, xlim=c (0,xmax),ylim=c (0, 1),
plot(x_vec,r_ac_vec, type="l", lwd=2, xlim=c (0,xmax),ylim=c (0, 1),
xlab=" Heightuof plantu(cm)", ylab="Probability", cex.lab=scale
xlab=" Heightuof plantu(cm)", ylab="Probability", cex.lab=scale
_text,cex.axis=scale_text)
_text,cex.axis=scale_text)
points(dat_repr$h, dat_repr$reproduced, pch="|", cex=line_frac)
points(dat_repr$h, dat_repr$reproduced, pch="|", cex=line_frac)
mtext("(d)", side=3,adj=dx, padj=dy, line= -1.3)
mtext("(d)", side=3,adj=dx, padj=dy, line= -1.3)
par(mfrow=c (1,1))

```
par(mfrow=c (1,1))
```


## B．5．2 Plot recruit size distribution

```
###
## - Plot recruit size distn
# Specify data set to use
dat <- seedlingdata
dat_recr <- dat$Height2003
# Vectors to help plot f_e
ns_vec <- seq (0,10000, length=1000)
fe_vec <- expected_seedlings(ns_vec)
# Vectors to help plot size distribution
xmin <- min(dat_recr)
```

```
xmax \(<-\) max (dat_recr)
\(x_{-}\)vec \(<-\)seq (xmin, xmax, length=100)
y_vec <- recruit_size (x_vec)
\# PLOTS
\(\operatorname{par}(m f r o w=c(1,2))\)
distnce <- 0.02
scale _text \(<-1\)
plot (ns_vec, fe_vec, type="l", lwd=2,
```



```
    cex. lab=scale_text, cex.axis=scale_text, main="")
mtext (" (a)", side \(=3\), adj=distnce, \(\operatorname{line}=-1.3\) )
hist (dat_recr, breaks \(=7, x \lim =c(x \min -5, x m a x), y \lim =c(0,0.021)\),
```



```
    cex. lab=scale_text, cex.axis=scale_text, main="", freq=FALSE)
lines (x_vec, y_vec, type="l", lwd=2)
mtext (" (b)", side \(=3\), adj=distnce, line \(=-1.3\) )
\(\operatorname{par}(\operatorname{mfrow}=c(1,1))\)
```


## B. 6 Results Plots

## B.6.1 Simulate for 30 years

```
###
## -- Run simulations
# Run simulation starting from equilibrium
L<-5; U <- 550; m <- U-L
num_steps <- 30
# With seeds
ns_init <- 1
nt_init <- rep(0,m)
```

```
nr_init <- rep(0,m)
sim_res_rl <- master_local_simulation_ws(L,U,m, num_steps, ns_
    init,nt_init,nr_init, "rl")
sim_res_ac<- master_local_simulation_ws(L,U,m, num_steps, ns_
    init,nt_init,nr_init, "ac")
###
## - Seed production results
# Number of years in which adult competition model predicts more
    seeds
year_overtook <- 0
while (sim_res_ac$seeds[year_overtook+1] >= sim_res_rl$seeds[year
        _overtook+1]) {
    year_overtook <- year_overtook+1
}
year_overtook
years_more <- year_overtook-1
# Comparison of asymptotic behaviour
lower_seed <- min(tail(sim_res_ac$seeds,2))
upper_seed <- max(tail(sim_res_ac$seeds,2))
vs_rl <- 1-upper_seed/max(sim_res_rl$seeds)
lower_seed
upper_seed
vs_rl
# Plot initial seed production and long-term seed production
plot (0:num_steps, sim_res_rl$seeds, col="black",pch=20,
    xlab="Year", ylab="Numberьof seeds")
points(0:num_steps, sim_res_ac$seeds, col="black", pch=18)
```


## B.6.2 Mean height comparison

```
###
## - Mean height calculations
# Observed mean height of recruits, vegetative plants, and
    flowering plants
mean(seedlingdata$h1)
mean(hmdata_adult $h1)
mean(hmdata_flow $h1)
# Mean of recruits predicted by recruit size dist'n
h <- (U-L)/m; meshpts <- L + (1:m)*h - h/2
meshpts %*% recruit_size(meshpts)
# Mean of adult plants predicted at equilibrium
weighted_distn <- sim_res_rl$adults[num_steps,]/sum(sim_res_rl$
    adults[num_steps,])
meshpts %*% weighted_distn
# Mean of flowering plants predicted at equilibrium
weighted_distn <- sim_res_rl$flowering[num_steps,]/sum(sim_res_rl
    $flowering[num_steps,])
meshpts %*%% weighted_distn
```

